

Quantile Risk Management in Equity-Linked Life Insurance with Stochastic Interest Rate

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Insurance company can issue a *mixed* contract for the period [0,T], where the payoff function is a function of stock prices  $S_0, ..., S_T$  and T(x)

(future lifetime of a client of age x).

#### □ Terminology

- Segregated funds (Canada),
- Variable annuities (USA),
- Equity (unit)-linked insurance (Europe).

#### Numbers

According The Insured Retirement Institute (IRI, formerly, NAVA): "Variable annuity sales for the 3-rd quarter were \$34 billion. Year-to year quarterly sales of variable annuities were up, posting a 9.7% increase from third quarter 2009 sales".



#### **Types and Features of Contracts**

- > Death and maturity guarantees.
- > Term of the guarantee.
- Resets (policyholder is allowed to reset the guarantee at the current fund value).
- Fund switching (policyholder has the right to switch investments between various funds).

#### **Types of Hedging**

> *Perfect* hedging:

$$P\{X_T^{\pi^*} \ge H\} = 1$$

*Mean-variance* hedging:

 $E(X_T^{\pi^*} - H)^2$  is minimal



 $\succ$ 

$$E\{l(H-X_T^{\pi^*})^+\}$$
 is minimal

where l is some loss function.

In particular, for *quantile* hedging, we have  $l(x) = I_{(0,\infty)}(x)$  and  $P\{X_T^{\pi^*} \ge H\}$  is maximized.





- Using Hedging for Pricing and Risk Management: construct a strategy that exactly replicates the cash flows of a contingent claim
- Exact replication is not possible: find a strategy with a cash flow "close enough" to the payoff of the contingent claim in some probabilistic sense
- Equity-linked life insurance contracts have a mortality component => the exact replication is not possible

# Introduction References on Equity-linked Life Insurance

- □ Brennan and Schwartz (1976, 1979)
- □ Boyle and Schwartz (1977)
- □ Bacinello and Ortu (1993)
- □ Aase and Person (1994)
- □ Ekern and Person (1996)
- □ Moeller (1998, 2001)
  - Contracts with fixed or deterministic guarantees
  - Reduced them to call/put options
  - Apply perfect or mean-variance hedging to calculate prices

# Introduction References on Equity-linked Life Insurance

- $\square \quad \text{Melnikov} (2004)$
- □ Melnikov and Skornyakova (2005)
  - > Contracts with stochastic guarantee
  - > Reduce them to embedded option  $\left(S_T^1 S_T^2\right)^+$
  - > Apply quantile hedging to calculate prices
- □ Quansheng Gao, et.al (2010)
  - Guarantee of the contracts grows exponentially over time
  - Stochastic interest rate
  - Diffusion model

## Introduction Contents



- □ Briefly review the work of Melnikov and Skornyakov (2005)
- Consider a two-factor jump-diffusion model to describe a financial market and HJM framework for stochastic interest rate
- □ Study equity-linked pure endowment contracts with a stochastic guarantee
- □ Use quantile hedging technique for pricing these contracts
- □ Illustrate our results with actual data

# Review: Financial settings Melnikov and Skornyakova (2005)



- > Non-risky asset  $B_t = \exp(rt), t \ge 0, r \ge 0$
- ► Risky assets  $S_t^1$  and  $S_t^2$  on  $(\Omega, \mathbf{F}, P)$ , prices follow the jump-diffusion model  $dS_t^i = S_{t-}^i(\mu_i dt + \sigma_i dW_t - \nu_i d\Pi_t), i = 1,2$
- $\square \quad \text{Market is complete, the unique risk-neutral probability has the density} \\ Z_T = \exp\left\{\alpha^* W_T \frac{(\alpha^*)^2}{2}T + (\lambda \lambda^*)t + \ln\frac{\lambda^*}{\lambda} \cdot \Pi_t\right\} \\ (\alpha^*, \lambda^*) \text{ are the unique solutions to } \begin{cases} \mu_1 r = -\sigma_1 \alpha^* + v_1 \lambda^*, \\ \mu_2 r = -\sigma_2 \alpha^* + v_2 \lambda^*, \ \lambda^* > 0 \end{cases}$
- **Conditions:**

$$\mu_{1} > \mu_{2}, \sigma_{1} > \sigma_{2}, \nu_{1}, \nu_{2} < 1, \quad \begin{cases} \sigma_{2}\nu_{1} - \sigma_{1}\nu_{2} \neq 0, \\ (\mu_{1} - r)\sigma_{2} - (\mu_{2} - r)\sigma_{1} \\ \sigma_{2}\nu_{1} - \sigma_{1}\nu_{2} \end{cases} > 0$$

# Review: Insurance settings Melnikov and Skornyakova (2005)



 $\Box$  T(x) on  $(\widetilde{\Omega}, \widetilde{\mathbf{F}}, \widetilde{P})$  is the remaining life time of a person at age x

 $\square \quad _T p_x = \widetilde{P}\{T(x) > T\} \text{ is a survival probability}$ 

 $\square$  Assumption:  $(\Omega, \mathbf{F}, P)$  and  $(\widetilde{\Omega}, \widetilde{\mathbf{F}}, \widetilde{P})$  are independent

□ Mortality risk arises from the dependence of the payoff on the survival status of a client at maturity  $H(T(x)) = H \cdot I_{\{T(x) > T\}}$ 

 $\square \quad \text{The payoff for the model:} \quad H(T(x)) = \max\left\{S_T^1, S_T^2\right\} \cdot I_{\{T(x) > T\}}$ 

# Main Results: Theorem



- Financial market is described by two-factor jump-diffusion model
- Equity-linked life insurance contract with a flexible guarantee

• On a set 
$$\{\Pi_T = n\}$$
, if

$$\frac{(\mu_1 - r)\nu_1 - (\mu_2 - r)\nu_2}{\sigma_2 \nu_1 - \sigma_1 \nu_2} = \frac{\sigma_1 \sigma_2}{\sigma_2 + (\sigma_1 - \sigma_2)^2}$$

Then

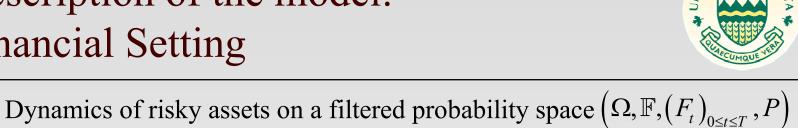
Then  

$$\sum_{T}^{\infty} p_{n,T}^{*} \left( \upsilon_{n,T}^{1} \Phi \left( b_{+} \left( \upsilon_{n,T}^{1}, d_{n} \upsilon_{n,T}^{2}, T \right) \right) - \upsilon_{n,T}^{2} \Phi \left( b_{-} \left( \upsilon_{n,T}^{1}, d_{n} \upsilon_{n,T}^{2}, T \right) \right) \right)$$

$$\sum_{n=0}^{\infty} p_{n,T}^{*} \left( \upsilon_{n,T}^{1} \Phi \left( b_{+} \left( \upsilon_{n,T}^{1}, \upsilon_{n,T}^{2}, T \right) \right) - \upsilon_{n,T}^{2} \Phi \left( b_{-} \left( \upsilon_{n,T}^{1}, \upsilon_{n,T}^{2}, T \right) \right) \right)$$
where  

$$\upsilon_{n,T}^{i} = (1 - \nu_{i})^{n} e^{\nu_{i}}, i = 1, 2, \qquad p_{n,T}^{*} = \exp\left\{ -\lambda^{*}T \right\} \frac{\left( \lambda^{*}T \right)^{n}}{n!}$$

#### Description of the model: **Financial Setting**



$$dS_t^i = S_{t-}^i \left( \mu_i dt + \sigma_i dW_t - \nu_i d\pi_t \right)$$

where  $\mu_i \in R, \sigma_i > 0, \nu_i < 1$  and  $\sigma_1 > \sigma_2$ 

HJM framework > f(t,T) : instantaneous forward interest rate  $f(t,T) = f(0,T) + \int_0^t \alpha(s,T) ds + \int_0^t \sigma(s,T) dW_s + \int_0^t \beta \left[ d\Pi_s - \lambda ds \right]$ > r(t) : spot interest rate  $r(t) = f(t,t) = f(0,t) + \int_0^t \alpha(s,t) ds + \int_0^t \sigma(s,t) dW_s + \int_0^t \beta[d\Pi_s - \lambda ds]$ 

# Description of the model: Financial Setting



- > P(t,T): price of a default-free discount zero-coupon bond at time t with maturity time T  $P(t.T) = \exp\left(-\int_{t}^{T} f(t,s) ds\right)$ > B(t): accumulated money account  $B(t) = \exp\left(\int_{0}^{t} r(s) ds\right)$
- □ Market is complete if

$$\frac{\mu_1 - r(t) \sigma_2 - (\mu_2 - r(t)) \sigma_1}{\sigma_2 v_1 - \sigma_1 v_2} > 0 \quad \text{and} \quad \sigma_2 v_1 - \sigma_1 v_2 \neq 0$$

• The unique martingale measure  $P^*$  has the local density

$$Z_{t} = \frac{dP^{*}}{dP}\Big|_{t} = \exp\left\{\int_{0}^{t}\phi_{s}dW_{s} - \frac{1}{2}\int_{0}^{t}\phi_{s}^{2}ds + \int_{0}^{t}(\lambda - \lambda_{s}^{*})ds + (\ln\lambda_{t}^{*} - \ln\lambda)\Pi_{t}\right\}$$
  
$$\left(\phi_{t}, \lambda_{t}^{*}\right) \text{ are the solutions for the equations } \begin{cases}\mu_{1} - r(t) + \phi_{t}\sigma_{1} - \nu_{1}\lambda_{t}^{*} = 0\\\mu_{2} - r(t) + \phi_{t}\sigma_{2} - \nu_{2}\lambda_{t}^{*} = 0\end{cases}$$

## Description of the model: Financial Setting



Explicit representations of B(t) and  $S_t^i$  in terms of the parameters of the system:

$$B(t) = \frac{1}{P(0,t)} \exp\left\{\frac{1}{2} \int_{0}^{t} (\sigma^{*}(s,T))^{2} ds + \int_{0}^{t} \left[e^{-\beta(T-s)} - 1\right] \lambda_{s}^{*} ds + \int_{0}^{t} \sigma^{*}(s,T) dW_{s}^{*} - \int_{0}^{t} \beta(T-s) d\Pi_{s}\right\},$$
  
$$S_{t}^{i} = S_{0}^{i} B(t) \exp\left\{\sigma_{i} W_{t}^{*} + \Pi_{t} \ln(1-v_{i}) + \int_{0}^{t} \left(v_{i} \lambda_{s}^{*} - \frac{1}{2} \sigma_{i}^{2}\right) ds\right\}$$

## Description of the model: Insurance Setting



- $\square$  Assumption:  $(\Omega, \mathbb{F}, P)$  and  $(\tilde{\Omega}, \tilde{F}, \tilde{P})$  are independent.
- □ A single premium equity-linked life insurance contract with payoff  $C(t) = \max(S_t^1, S_t^2)$ 
  - $S^1$ : maximal size of future profit
  - $S^2$ : stochastic guarantee for the insured
- □ The initial price of the contract: (Brennan-Schwartz price)

 $H(0) = E^* \left\{ \tilde{E} \left[ C(T) B^{-1}(T) I \{ T_x > T \} \right] \right\} = E^* \left[ C(T) B^{-1}(T) \right] \cdot_T p_x$ 

- □ Perfect hedging is not possible due to a budget constraint  $H(0) < E^*[C(T) \cdot B^{-1}(T)]$
- □ Find a strategy that will hedge successfully with the maximal probability

## Quantile Hedging Definitions



□ Self-financing strategy  $\pi$  has a budget constraint

$$H_0^{\pi} \le H_0 < E^* C_T e^{-rT}$$

 $\Box \qquad A\!\!\left(H_0^{\pi},\pi\right) = \!\left\{\omega: H_T^{\pi}\!\left(H_0^{\pi}\right) \ge C_T\right\} \quad \text{is a successful hedging set}$ 

 $\square$   $\pi^*$  is a *quantile hedge* if

$$P\{A(H_0^{\pi^*}, \pi^*)\} = \max_{\pi: H_0^{\pi} \le H_0} P\{A(H_0^{\pi}, \pi)\}$$

□ How to construct the quantile hedge  $\pi^*$  and the successful hedging set  $A(H_0^{\pi^*}, \pi^*)$ ?

# Quantile Hedging Methodology



The answer is given in the following fact (Foellmer and Leukert (1999)) Let  $A^* \in F_T$  be a solution to the problem  $P(A^*) = \max_{A \in F_T : E^*(C_T : I_A) \leq H_0} P(A)$ 

Then the quantile hedge  $\pi^*$ 

- > does exist
- > is unique
- > is a perfect hedge for a modified claim  $C_{A^*} = C_T \cdot I_{A^*}$

The structure of a maximal successful hedging set is  $A^* = \left\{ Z_T^{-1} > a \frac{C_T}{B_T} \right\}$ , where A constant *a* is defined  $a = \inf \left\{ c : P^* \left( \frac{dP}{dP^*} > c \right) \le \alpha \right\}$ 

# Equity-Linked Life Insurance Connecting Financial and Insurance Risks



Due to the structure of the price,  $H(0) = E^* [C(T)B^{-1}(T)] \cdot {}_T p_x$ 

we apply quantile hedging, 
$$H(0) = E^* \left[ C(T) I_{\{A^*\}} \cdot B^{-1}(T) \right]$$

□ Key formulae connecting financial and insurance risks

$${}_{T} p_{x} = \frac{E^{*} \left[ C(T) I_{\{A^{*}\}} \cdot B^{-1}(T) \right]}{E^{*} \left[ C(T) \cdot B^{-1}(T) \right]}$$

# Application of Quantile Hedging Preliminary Calculations

Maximal successful hedging sets

$$A^{*} = \left\{ Z_{T}^{-1} \ge a_{n} \frac{C_{T}}{B(T)} \right\} = \left\{ Z_{T}^{-1} \ge a_{n} \frac{\max\left(S_{T}^{1}, S_{T}^{2}\right)}{B(T)} \right\}$$

$$\Box \quad \text{The price of the contract } H(0) = E^* \left[ \max\left(S_T^1, S_T^2\right) B^{-1}(T) I_{\{A^*\}} \right] \\ = E^* \left[ \frac{S_T^1}{B_T} I_{\{A^*\}} \cdot I_{\{S_T^1 > S_T^2\}} \right] + E^* \left[ \frac{S_T^2}{B_T} I_{\{A^*\}} \cdot I_{\{S_T^1 \le S_T^2\}} \right]$$

□ Further analysis relies on change of measure technique and properties of jump-diffusion processes



## Main Results Theorem 1



- Financial market is described by two-factor jump-diffusion model
- Stochastic interest rate is in HJM framework
- Equity-linked life insurance contract with a flexible guarantee

Then the Brennan-Schwartz price of the contract is

$$H(0) = \sum_{n=0}^{\infty} p_{n,T}^{*} \left[ e^{v_{1} \int_{0}^{T} \lambda_{t}^{*} dt} (1-v_{1})^{n} S_{0}^{1} \Psi_{1}(\Gamma_{1},\Gamma_{2}) + e^{v_{2} \int_{0}^{T} \lambda_{t}^{*} dt} (1-v_{2})^{n} S_{0}^{2} \Psi_{2}(\tilde{\Gamma}_{1},\tilde{\Gamma}_{2}) \right]$$

$$P_{n,T}^{*} = e^{-\int_{0}^{T} \lambda_{t}^{*} dt} \frac{\left(\int_{0}^{T} \lambda_{t}^{*} dt\right)^{n}}{n!}$$

where

# Main Results Theorem 1



$$\begin{split} \Gamma_{1} &= -\ln \frac{a_{n} \cdot S_{0}^{1} \left(\lambda_{T}^{1}\right)^{n}}{\lambda^{n}} - \frac{1}{2} \delta_{1}^{2} - \int_{0}^{T} \left(\lambda - \lambda_{s}^{1}\right) ds \\ \Gamma_{2} &= \ln \frac{S_{0}^{1} \left(1 - v_{1}\right)^{n}}{S_{0}^{2} \left(1 - v_{2}\right)^{n}} - \int_{0}^{T} \lambda_{s}^{*} \left(v_{2} - v_{1}\right) ds + \frac{1}{2} \delta_{2}^{2} \\ \tilde{\Gamma}_{1} &= -\ln \frac{a_{n} \cdot S_{0}^{2} \left(\lambda_{T}^{2}\right)^{n}}{\lambda^{n}} - \frac{1}{2} \tilde{\delta}_{1}^{2} - \int_{0}^{T} \left(\lambda - \lambda_{s}^{2}\right) ds \\ \tilde{\Gamma}_{2} &= \ln \frac{S_{0}^{2} \left(1 - v_{2}\right)^{n}}{S_{0}^{1} \left(1 - v_{1}\right)^{n}} - \int_{0}^{T} \lambda_{s}^{*} \left(v_{1} - v_{2}\right) ds + \frac{1}{2} \tilde{\delta}_{2}^{2} \\ \lambda_{t}^{1} &= \lambda_{t}^{*} \left(1 - v_{1}\right) \qquad \lambda_{t}^{2} &= \lambda_{t}^{*} \left(1 - v_{2}\right) \end{split}$$

# Main Results Theorem 2



The survival probability of an insured is as following:

$${}_{T}p_{x} = \frac{\sum_{n=0}^{\infty} p_{n,T}^{*} \left[ p\Psi(\Gamma_{1},\Gamma_{2};\rho,\delta_{1},\delta_{2}) + q\Psi(\tilde{\Gamma}_{1},\tilde{\Gamma}_{2};\tilde{\rho},\tilde{\delta}_{1},\tilde{\delta}_{2}) \right]}{\sum_{n=0}^{\infty} p_{n,T}^{*} \left[ p\Phi\left(\frac{d_{1}}{(\sigma_{1}-\sigma_{2})\sqrt{T}}\right) + q\Phi\left(\frac{d_{2}}{(\sigma_{1}-\sigma_{2})\sqrt{T}}\right) \right]}$$

Where 
$$d_1 = \ln \frac{S_0^1 (1 - v_1)^n}{S_0^2 (1 - v_2)^n} - \int_0^T \lambda_s^* (v_2 - v_1) ds + \frac{1}{2} \delta_2^2$$
,  $p_n = e^{v_1 \int_0^T \lambda_t^* dt} (1 - v_1)^n S_0^1$ 

$$d_{2} = \ln \frac{S_{0}^{2} (1 - v_{2})^{n}}{S_{0}^{1} (1 - v_{1})^{n}} - \int_{0}^{T} \lambda_{s}^{*} (v_{1} - v_{2}) ds + \frac{1}{2} \tilde{\delta}_{2}^{2} \quad , \quad q_{n} = e^{v_{2} \int_{0}^{T} \lambda_{t}^{*} dt} (1 - v_{2})^{n} S_{0}^{2}$$

#### Remark: How to determine the constant *a*

□ Fix a probability  $\mathcal{E}$  of failure to hedge on each set  $\{\Pi_T = n\}$ Or equivalently, fix the probability of successful hedging as

Where  

$$P\left(A^* | \pi_T = n\right) = 1 - \varepsilon = \Phi\left(\Delta\right) \qquad \Delta = \Delta_1 \cup \Delta_2$$

$$\Delta_1 = \frac{-\ln \frac{a \cdot S_0^1 \left(\lambda_T^1\right)^n}{\lambda^n} - \frac{1}{2} \delta_1^2 - \int_0^T \left(\lambda - \lambda_s^1\right) ds}{\delta_1}$$

$$\Delta_2 = \frac{-\ln \frac{a \cdot S_0^2 \left(\lambda_T^2\right)^n}{\lambda^n} - \frac{1}{2} \tilde{\delta}_1^2 - \int_0^T \left(\lambda - \lambda_s^2\right) ds}{\tilde{\delta}_1}$$

 $\Box$  Using the log-normality of the conditional distribution to estimate a



# Numerical Illustration Inputs



- □ Contracts with flexible guarantee: Russell 2000 and the S&P 500
- □ Transfer to One factor Vasicek-Hull-White model
- Estimated parameters for two-factor jump diffusion model (from monthly observations from 09/1987 to 09/2010)

$$\mu_1 = 0.2763$$
  $\mu_2 = 0.2898$   $\sigma_2 = 0.15$   $\sigma_1 = 0.19$ 

$$v_1 = -0.27$$
  $v_2 = -0.2$   $\lambda = 0.17$ 

- $\Box$   $S_0 = 1000$  is an initial investment
- $\Box$  T = 1, 3, 5, 10, 15, 20 years are terms of the contracts
- **\Box** Fix different level of financial risks  $\varepsilon = 0.01, 0.025, 0.05$



# Numerical Illustration

#### **Table 1:** Survival probabilities $_T p_x$ with flexible guarantee

Т	ε =0.01	ε =0.025	ε =0.05
1	0.9885	0.9718	0.9447
3	0.9878	0.9705	0.9426
5	0.9874	0.9697	0.9413
10	0.9867	0.9684	0.9391
15	0.9859	0.9667	0.9364
20	0.9853	0.9656	0.9345



# Numerical Illustration

#### □ Table 2: Age of insured with flexible guarantee

Т	ε =0.01	ε =0.025	ε =0.05
1	58	69	78
3	45	55	63
5	39	48	56
10	23	39	46
15	12	31	39
20	6	24	33

# Numerical Illustration from Melnikov and Skornyakova (2005)



□ Table 1: Survival probabilities with flexible guarantee

Т	ε =0.01	ε =0.025	ε =0.05
1	0.9447	0.8774	0.7811
3	0.9511	0.8910	0.8041
5	0.9549	0.9108	0.8387
10	0.9605	0.9108	0.8378

□ Table 2: Age of Insured with flexible guarantee

Т	ε =0.01	ε =0.025	ε =0.05
1	78	87	94
3	61	71	79
5	53	63	71
10	41	50	58

## Numerical Illustration Conclusions



- □ Whenever the financial risk level  $\mathcal{E}$  increases, the survival probability<sub>T</sub>  $p_x$  decreases, so that the clients' age increases in the same period. The insurance company should attract older clients for the contract with flexible guarantee to compensate increasing financial risk.
- With longer contract maturities, the company is able to attract younger clients while maintaining the same financial risk exposure, as a survival probability is decreasing over time.
- □ For fixed contract maturity, as the financial risk  $\mathcal{E}$  increases, the change of survival probabilities  $_T p_x$  is not dramatic as the corresponding results in Melnikov and Skornyakova (2005).
- □ For fixed financial risk  $\varepsilon$ , as the contract maturity increases, the change of survival probabilities  $_T P_x$  decreases in small percentage while the results in Melnikov and Skornyakova (2005) shows the opposite way, which is increasing.



#### Further Developments Discussion

- Mortality Modeling
  - > Use theoretical models of mortality Gompertz, Makeham, Lee-Carter
  - > Allows to take into account new tendencies in mortality
- □ Modeling with other Risk Measures
  - Conditional Tail Expectation:

$$CTE_{1-\varepsilon}(\pi) = E\left(H - X_T^{\pi}(\pi) \middle| H - X_T^{\pi}(\pi) \ge VaR_{1-\varepsilon}(\pi)\right)$$

> Rockafellar and Uryasev (2002):

If 
$$\overline{z}$$
 is a solution to  $RU(z) = z + \frac{1}{\varepsilon} \left[ \inf_{\pi} E(H - X_T^{\pi}(x) - z)^+ \right] \rightarrow \min$ 

then  $VaR_{1-\varepsilon}(\pi^*) = \overline{z}$ ,  $CTE_{1-\varepsilon}(\pi^*) = RU(\overline{z})$ 



#### Further Developments Discussion

- □ Shortfall minimization problem for the claim  $H(z) = (H z)^+$   $E(H - X_T^{\pi}(x) - z)^+ = E((H - z)^+ - X_T^{\pi}(x))^+ \rightarrow \inf$ over all strategies with initial budget constraints
- $\square \quad \left(\widetilde{X}_0(z), \widetilde{\pi}(z)\right) \text{ solution to this problem (Foellmer and Leukert (2000)),} \\ \text{ where } \quad \widetilde{\pi}(z) \text{ is a perfect hedge to a modified claim}$

$$\widetilde{H}(z) = H(z)I_{\{Z_T^{-1} > \widetilde{a}(z)\}},$$
  
$$\widetilde{a}(z) = \inf\left(a \ge 0 : E^*H(z)I_{\{Z_T^{-1} > a\}} \le X_0\right)$$

The function 
$$RU(z)$$
 has the following structure  
 $RU(z) = z + \frac{1}{\varepsilon} EI_{\{Z_T^{-1} \le \widetilde{\alpha}(z)\}} H(z), \quad z < z^*, \quad RU(z) = z, \quad z \ge z^*$   
where  $E^*H(z^*) = X_0$ 



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