

Likelihood Adaptive Modified Penalty and its properties

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Outline

- 1 Introduction
- 2 Likelihood Adaptive Modified Penalty (LAMP)
- 3 Theoretical Properties
 - Asymptotic stability
 - Oracle Properties
- 4 Coordinate Descent Algorithm
- 5 Simulation
- 6 Summary and Future Work

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History of Penalty Functions

- L_0 penalty: AIC (Akaike, 1973), C_p (Mallows, 1973), BIC (Schwartz, 1978), RIC (Foster and George, 1994), Extended BIC (Chen and Chen, 2008), ...
- L_q penalty (bridge penalty): $0 < q \leq 2$, Frank and Friedman (1993)
- Non-negative garrote: Breiman (1995)
- L_1 penalty (LASSO): Tibshirani (1996), Osborne et. al (2000), Efron et. al (2004)
- SCAD penalty: Fan and Li (2001)
- Elastic Net: Zou and Hastie (2005)
- Adaptive LASSO: Zou (2006)
- Minimax concave penalty (MCP): Zhang (2010)
-
-

Sparsity and Stability

- Convex penalties: very stable, efficient algorithms.
- Non-convex penalties: gain sparsity, but sacrifice some stability.
- SCAD and MCP: choose an appropriate γ .

Goal

Good balance between
Sparsity and Stability.

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Negative Absolute Prior (NAP)

- X_1, \dots, X_n , i.i.d, with pdf $f(x, \beta)$
- **Negative Absolute Prior (NAP) for β :**
 $p(\beta) \propto \prod_{i=1}^k 1/f(C_i, |\beta|)$ for some integer $k > 0$ and k constants $C_i, i = 1, \dots, k$.
- The prior: remove noisy sample information by selecting appropriate C_i 's.
- $|\beta|$: treats unknown sign of β .
- A conjugate prior if $p(\beta) \propto \prod_{i=1}^k f(C_i, \beta)$.

Example

$\mathbf{X}_1, \dots, \mathbf{X}_n$, i.i.d, with pdf $f(\alpha + \mathbf{X}^T \beta)$. Then
 $\beta_j \propto 1/f(c_1 + c_2 |\beta_j|)$ for some constants c_1 and c_2 is a NAP for
 $\beta = (\beta_1, \dots, \beta_p)$.

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Likelihood Adaptive Modified Penalty (LAMP)

- For generalized linear model

$Y_i | \mathbf{X}_i; \boldsymbol{\theta} \sim i.i.d \exp(T(y)\xi - g(\xi) + h(x)), i = 1, \dots, n,$ where
 $\xi = \mathbf{X}^T \boldsymbol{\theta}$, and $\mathbf{X} = (1, X_1, X_2, \dots, X_p)^T$, $\boldsymbol{\theta} = (\alpha, \boldsymbol{\beta}^\tau)^\tau$.

- Log-likelihood:

$$l = \frac{1}{n} \sum_{i=1}^n [T(Y_i)\xi_i - g(\xi_i)],$$

- We propose to maximize:

$$l - \sum_{j=1}^p \frac{\lambda_n^2}{-g'(\alpha_1)\lambda_0} [\alpha_1 + \frac{\lambda_0}{\lambda_n} |\beta_j| - g(\alpha_1 + \frac{\lambda_0}{\lambda_n} |\beta_j|)].$$

- Taking a NAP*: bayesian interpretation.

Likelihood Adaptive Modified Penalty (LAMP)

- For convenience, take away $\frac{\lambda_0}{\lambda_n}|\beta_j|$, define
Likelihood Adaptive Modified Penalty:

$$p_{\lambda_n}(\beta_j) = \frac{\lambda_n^2}{g'(\alpha_1)\lambda_0} [g\left(\frac{\lambda_0}{\lambda_n}\beta_j + \alpha_1\right) - g(\alpha_1)], \quad (1)$$

where α_1 and λ_0 are positive constants.

- For identifiability purpose: $p_{\lambda_n}(0) = 0$ and $p'_{\lambda_n}(0) = \lambda_n$.
- Note that:

$$p'_{\lambda_n}(\theta) = \lambda_n g'\left(\frac{\lambda_0}{\lambda_n}\theta + \alpha_1\right)/g'(\alpha_1),$$

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Quick Facts

- For linear regression, $g(\xi) = \xi^2/2$, LAMP reduces to Elastic Net (Zou and Hastie, 2005);
- Smooth $g(\cdot)$, smooth LAMP in $(0, \infty)$.
- SCAD and MCP $\notin C^2(0, \infty)$.
- Intuition: smoothness in penalty \rightarrow more stable optimization problem.

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- Smooth $g(\cdot)$, smooth LAMP in $(0, \infty)$.
- SCAD and MCP $\notin \mathcal{C}^2(0, \infty)$.
- Intuition: smoothness in penalty \rightarrow more stable optimization problem.

Other LAMP Examples

- Logistic regression: $g(\xi) = \log(1 + e^\xi) - \xi$,
 $p_{\lambda_n}(\theta) = \frac{\lambda_n^2(1+\rho)}{\lambda_0} \log[\frac{(1+\rho)e^{\lambda_0/\lambda_n\theta}}{1+\rho e^{\lambda_0/\lambda_n\theta}}]$, sigmoid penalty;
- Poisson regression: $g(\xi) = e^{-\xi}$, poisson penalty;
- Gamma regression: $g(\xi) = -k \log(\xi)$;
- Inverse gaussian regression: $g(\xi) = -l\sqrt{\xi}$;
- Probit model: $g(\xi) = -\log(\Phi(\xi))$.

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Comparison with other penalties

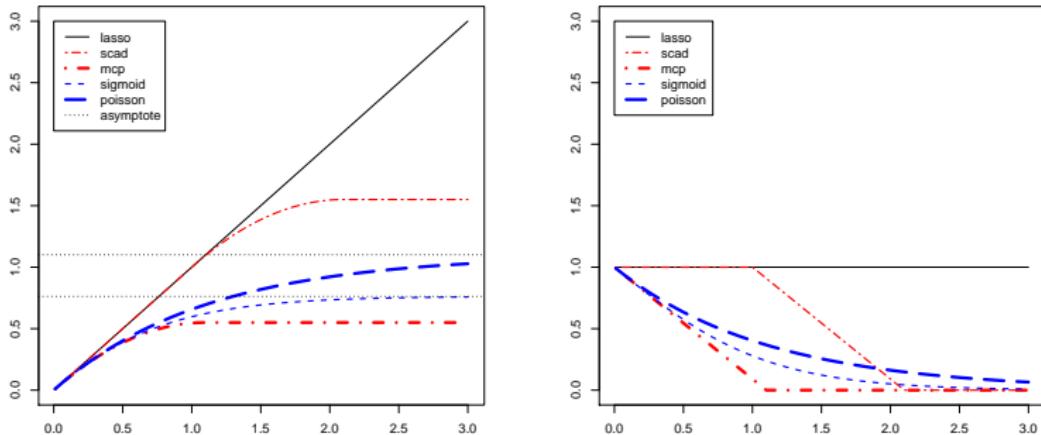


Figure: Penalty functions and their derivatives of the Lasso, SCAD, MCP, Sigmoid(for logistic), and Poisson. We choose $\lambda = 1$, $\gamma = 1.1$ for MCP, $\gamma = 2.1$ for SCAD, $\lambda_0 = 1/1.1$ for Sigmoid and Poisson penalty, and $\rho = 1$ for Sigmoid penalty. We choose parameters this way to keep the maximum concavity of these penalties the same.

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Asymptotic stability

Minimize random function

$$M_n(\boldsymbol{x}, \theta) = \frac{1}{n} \sum_{i=1}^n m(\boldsymbol{x}_i, \theta) + r_n(\theta), \theta \in \Theta.$$

Denote by $\operatorname{arglmin} M_n(\boldsymbol{x}, \theta)$

$$\{\theta^* \in \Theta \mid \text{for some } c > 0, M_n(\boldsymbol{x}, \theta^*) \leq M_n(\boldsymbol{x}, \theta), \forall \|\theta - \theta^*\| \leq c\}.$$

Two types of asymptotic stability

F_n denotes the domain for \mathbf{x} , $\text{diam}_d(A) \triangleq \max_{\mathbf{x}, \mathbf{y} \in A} \|\mathbf{x} - \mathbf{y}\|_d$ for measure d and set A .

We call $\text{arglmin}M_n(\mathbf{x}, \theta)$ satisfies **weak asymptotic stability** if
 $\forall \epsilon > 0, \forall \mathbf{x} \in F_n$ fixed,

$$P\left(\overline{\lim}_{n \rightarrow \infty} \overline{\lim}_{\delta \rightarrow 0} \text{diam}_d \bigcup_{\substack{\|\boldsymbol{\delta}\|_d < \delta \\ \mathbf{x} + \boldsymbol{\delta} \in F_n}} \{\text{arglmin}M_n(\mathbf{x} + \boldsymbol{\delta}, \theta)\} = 0\right) = 1; \quad (2)$$

Two types of asymptotic stability (Cont')

$\text{arglmin} M_n(\mathbf{x}, \theta)$ satisfies **strong asymptotic stability** if *weak asymptotic stability* holds and $\forall \epsilon > 0, \forall \mathbf{x} \in F_n$ fixed,

$$\varliminf_{n \rightarrow \infty} P\left(\overline{\lim_{\delta \rightarrow 0}} \text{diam}_d \bigcup_{\substack{\|\boldsymbol{\delta}\|_d < \delta \\ \mathbf{x} + \boldsymbol{\delta} \in F_n}} \{\text{arglmin} M_n(\mathbf{x} + \boldsymbol{\delta}, \theta)\} = 0\right) = 1. \quad (3)$$

Example of asymptotic stability

Let $\mathbf{x}_n^{(1)}$ and $\mathbf{x}_n^{(2)}$ be two local minimizers of $M_n(\mathbf{x}, \theta)$.

- Weak asymptotic stability:

$$P(\lim_{n \rightarrow \infty} \|\mathbf{x}_n^{(1)} - \mathbf{x}_n^{(2)}\|_d = 0) = 1.$$

- Strong asymptotic stability:

$$P(\lim_{n \rightarrow \infty} \mathbf{x}_n^{(1)} = \mathbf{x}_n^{(2)}) = 1.$$

Asymptotic stability

Lemma

Suppose $\forall \mathbf{x} \in F_n$. $m_n(\mathbf{x}, \theta) = \frac{1}{n} \sum_{i=1}^n m(\mathbf{x}_i, \theta)$ is a smooth strictly convex function of θ . \forall fixed θ , $m_n(\mathbf{x}, \theta)$ is continuous in \mathbf{x} .

- ① $r_n(\theta) > 0, r_n(\theta) \rightarrow 0$ uniformly in $\theta \in \Theta$ as $n \rightarrow \infty$,
- ② Each $\mathbf{x} \in F_n$ is a limit point.
- ③ $m_n(\mathbf{x}, \theta)$ is an open map: $F_n \rightarrow [\min(m_n(\mathbf{x}, \theta)), +\infty)$ with θ fixed,
- ④ $EM_n(\mathbf{x}, \theta)$ is strictly convex at $\theta_0 = \arg \min EM_n(\mathbf{x}, \theta)$.

$\operatorname{arglmin} M_n(\mathbf{x}, \theta)$ has weak asymptotic stability If (1)-(3) hold;
strong asymptotic stability if (1)-(4) hold.

Oracle Properties for LAMP

Theorem

We assume p is fixed. Let $\{\hat{\beta}\}$ be the set of local maximizers of $\tilde{l}(\beta) = \sum_{i=1}^n [T(y_i)\xi_i - g(\xi_i)] - n \sum_{j=1}^p p_{\lambda_n}(|\beta_j|)$, where $\xi_i = \tilde{X}_i^T \theta$. Under some regularity conditions,

- any $\hat{\beta}$ has estimation consistency and oracle property (in the sense defined by Fan and Li, 2001),
- $\{\hat{\beta}\}$ has strong asymptotic stability.

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Algorithm

We use the popular coordinate descent algorithm (Friedman et. al, 2010; Keerthi and Shevade, 2007) to calculate the solution path.

- ➊ Start with λ_{\max} where all the estimates are 0.
- ➋ Do optimization on a grid

$\tau\lambda_{\max} = \lambda_K < \lambda_{K-1} < \dots < \lambda_1 = \lambda_{\max}$. When $\lambda = \lambda_{i+1}$, use the solution from $\lambda = \lambda_i$ as the initial value to speed up. Here $\tau = 0.01$, $K = 100$.

- ➌ Inside each λ_i , we cycle through all the variables and update their coefficients until converge.

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- ③ Inside each λ_i , we cycle through all the variables and update their coefficients until converge.

Remarks on the algorithm

- In Step 3, when updating each variable, use Newton-Raphson algorithm.
- In Step 3, we calculate a *violation* measure for all the variables and update the variables according to the descending order of the violation (Keerthi and Shevade, 2007).

$$\text{viol}_j = \begin{cases} |F_j|, & \text{if } j = 0; \\ \max\{0, -n\lambda_n - F_j, -n\lambda_n + F_j\}, & \text{if } \beta_j = 0, j > 0; \\ |F_j - n\text{sgn}(\beta_j)p'_{\lambda_n}(|\beta_j|)|, & \text{if } \beta_j \neq 0, j > 0. \end{cases}$$

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Logistic Regression

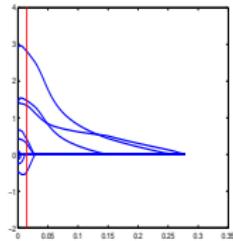
$n = 200$, logistic regression model with intercept $\alpha = -2$ and $\beta = (3, 2, 1, 0, 0, 0, 0, 0, 0, 0)^T$. $X_i \sim N(0, 1)$ and $\text{Corr}(X_i, X_j) = 0.5^{|i-j|}$. MRME stands for median of ratios of ME of a selected model to that of the ordinary least squares estimate under the oracle model. ME here stands for "model error" defined as:

$$\text{ME}(\hat{\theta}) = E[E_{\hat{\theta}}(Y|\mathbf{X}) - E_{\theta_0}(Y|\mathbf{X})]^2.$$

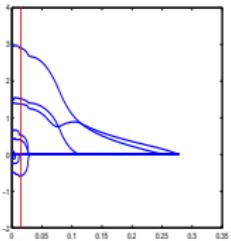
Logistic Regression (Cont')

| Penalties | TP | FP | cf | of | uf | L1 | L2 | MRME |
|-----------|------|------|------|------|------|------|------|------|
| sig(.08) | 2.67 | 0.56 | 0.37 | 0.30 | 0.33 | 2.28 | 1.22 | 0.51 |
| sig(.09) | 2.72 | 0.59 | 0.47 | 0.25 | 0.28 | 2.37 | 1.24 | 0.49 |
| sig(.1) | 2.69 | 0.55 | 0.42 | 0.27 | 0.31 | 2.43 | 1.26 | 0.50 |
| lasso | 2.99 | 3.25 | 0.04 | 0.95 | 0.01 | 2.54 | 1.18 | 0.50 |
| alasso(3) | 2.82 | 0.98 | 0.37 | 0.45 | 0.18 | 2.27 | 1.15 | 0.47 |
| alasso(4) | 2.69 | 0.52 | 0.44 | 0.25 | 0.31 | 1.92 | 1.05 | 0.41 |
| alasso(5) | 2.50 | 0.15 | 0.43 | 0.08 | 0.49 | 1.86 | 1.08 | 0.52 |
| mcp(2.8) | 2.79 | 0.66 | 0.37 | 0.42 | 0.21 | 2.05 | 1.13 | 0.48 |
| mcp(2.9) | 2.79 | 0.68 | 0.41 | 0.38 | 0.21 | 2.21 | 1.17 | 0.54 |
| mcp(3) | 2.81 | 0.87 | 0.36 | 0.45 | 0.19 | 2.33 | 1.21 | 0.50 |
| scad(2.4) | 2.80 | 1.01 | 0.25 | 0.55 | 0.20 | 2.25 | 1.20 | 0.54 |
| scad(2.7) | 2.87 | 1.15 | 0.30 | 0.57 | 0.13 | 2.43 | 1.27 | 0.55 |
| scad(3) | 2.89 | 1.24 | 0.24 | 0.65 | 0.11 | 2.20 | 1.17 | 0.50 |

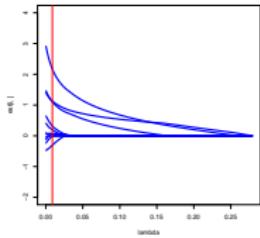
Solution Paths



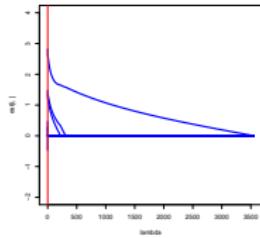
(a) Sigmoid (0.05)



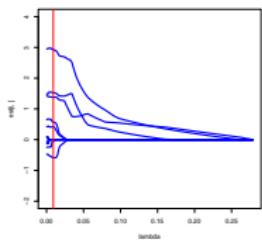
(b) Sigmoid (0.09)



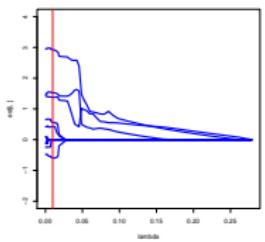
(c) Lasso



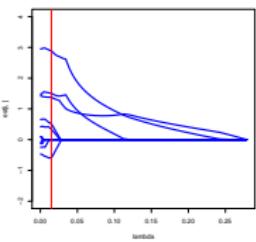
(d) ALasso (4)



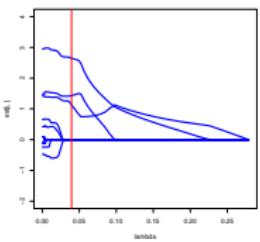
(e) SCAD (3)



(f) SCAD (2.7)



(g) MCP (2.9)



(h) MCP (2)

Figure: Solution Paths for all the penalties using different λ_0 or γ for selection.

Corresponding penalties

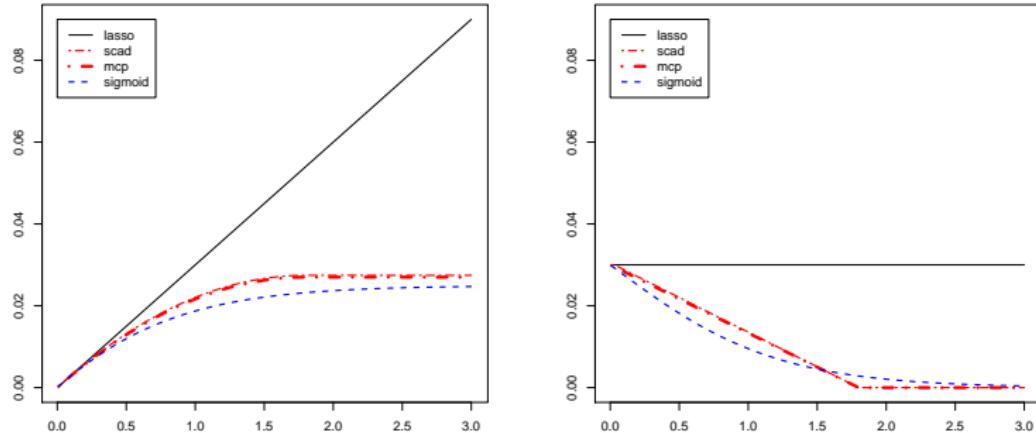


Figure: Left Panel: penalty functions. Right Panel: the first derivative of the penalties

Test the stability

Test stability of different penalties on logistic regression.

- ① 10-fold cross-validation → mean standard error of the estimates for the coefficients.
- ② Perturb the sample with gaussian error and do the previous simulation.
- ③ Repeat each simulation for 100 times

Test the stability (Cont')

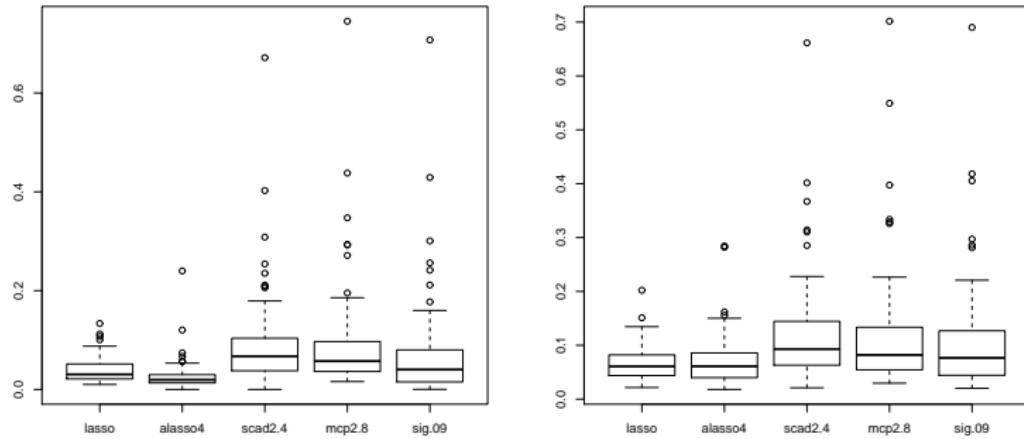


Figure: Boxplots for the mean standard error. Left: without random error.
Right: with random error.

Stability and Sparsity Balance: Revisited

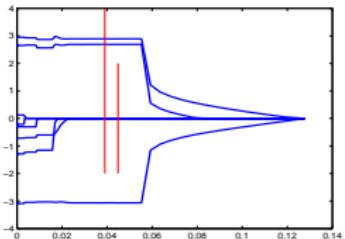
We use the same logistic regression model but with

$$\alpha = -4, \beta = (-2, 1.5, 0, 0, 2, 0, 0, 0)^T.$$

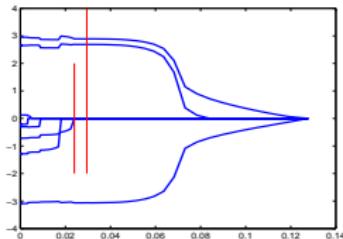
| Penalty | λ_0 | False0 | True0 | Penalty | γ | False0 | True0 |
|---------|-------------|--------|-------------|---------|----------|--------|-------------|
| Sigmoid | .04 | 0 | 4.69 | MCP | 1.1 | 0 | 4.89 |
| Sigmoid | .05 | .01 | 4.79 | MCP | 4 | 0 | 4.89 |
| Sigmoid | .06 | .01 | 4.85 | MCP | 12 | 0 | 4.89 |
| Sigmoid | .07 | .01 | 4.86 | MCP | 27 | 0 | 4.89 |
| Sigmoid | .09 | .01 | 4.89 | MCP | 34 | .01 | 4.90 |
| Sigmoid | .10 | 0 | 4.89 | MCP | 41 | .01 | 4.88 |
| Sigmoid | .12 | 0 | 4.89 | MCP | 48 | 0 | 4.83 |
| Sigmoid | .14 | 0 | 4.89 | MCP | 55 | 0 | 4.77 |
| Sigmoid | .16 | 0 | 4.89 | MCP | 62 | 0 | 4.71 |
| Sigmoid | .18 | 0 | 4.89 | MCP | 69 | 0 | 4.65 |

Table: BIC: 100 simulations

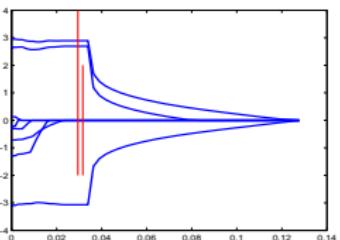
Solution Path Comparison



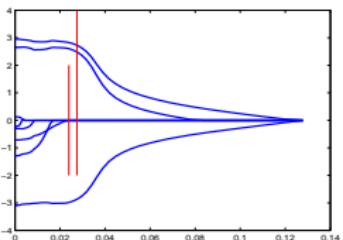
(a) MCP(34)



(b) Sigmoid(0.1)



(c) MCP(69)



(d) Sigmoid(0.04)

Figure: λ_0 and γ is chosen to make the penalties have similar sparsity.

Poisson Regression

$Y|\mathbf{X} \sim \text{Poisson}(\mathbf{X}^T \boldsymbol{\beta} + \alpha)$, \mathbf{X} follows the same distribution and correlation structure as before, and the true parameters $\boldsymbol{\beta} = (.6, .4, 0, 0, 1, 0, 0, 0, 0, 0, 0, 0, 0, 0)^T$, $\alpha = -1$. We perform 100 simulations with sample size $n = 100$.

Poisson Regression (Cont')

| Penalties | TP | FP | cf | of | uf | L1 | L2 | MRME |
|-----------|------|------|------|------|------|------|------|------|
| poi(2.7) | 2.82 | 0.47 | 0.53 | 0.29 | 0.18 | 0.61 | 0.34 | 0.21 |
| poi(3) | 2.80 | 0.52 | 0.56 | 0.24 | 0.20 | 0.65 | 0.36 | 0.22 |
| poi(10) | 2.80 | 0.61 | 0.51 | 0.29 | 0.20 | 0.66 | 0.36 | 0.23 |
| lasso | 3.00 | 4.05 | 0.04 | 0.96 | 0.00 | 1.01 | 0.43 | 0.37 |
| alasso(4) | 2.69 | 1.24 | 0.29 | 0.41 | 0.30 | 0.84 | 0.43 | 0.31 |
| alasso(5) | 2.52 | 0.64 | 0.31 | 0.27 | 0.42 | 0.80 | 0.44 | 0.28 |
| alasso(6) | 2.25 | 0.44 | 0.25 | 0.16 | 0.59 | 0.88 | 0.50 | 0.34 |
| mcp(1.1) | 2.81 | 0.84 | 0.46 | 0.35 | 0.19 | 0.74 | 0.38 | 0.30 |
| mcp(1.4) | 2.80 | 0.48 | 0.52 | 0.28 | 0.20 | 0.64 | 0.35 | 0.23 |
| mcp(1.7) | 2.80 | 0.83 | 0.47 | 0.33 | 0.20 | 0.73 | 0.38 | 0.28 |

Outline

1 Introduction

2 Likelihood Adaptive Modified Penalty (LAMP)

3 Theoretical Properties

- Asymptotic stability
- Oracle Properties

4 Coordinate Descent Algorithm

5 Simulation

6 Summary and Future Work

Summary

- Motivation: balance between stability and sparsity
- Solution: Likelihood Adaptive Modified Penalty (LAMP)
- Conclusion: strong asymptotic stability and oracle properties.

Future Work

- Diverging p .
- Characterization of degrees of stability.

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