



Density Models for default risk

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In many models for credit risk, it is implicitly assumed that the intensity contains all the needed information. Our goal is to present a more general setting.

Gapeev, P. V., Jeanblanc, M., Li, L., and Rutkowski, M. (2009):
Constructing Random Times with Given Survival Processes and Applications to
Valuation of Credit Derivatives. Forthcoming in: *Contemporary Quantitative
Finance*, C. Chiarella and A. Novikov, eds., Springer-Verlag, Berlin Heidelberg New
York, 2010.

Jeanblanc, M. and Song, S. (2010) Explicit Model of Default Time with given
Survival Probability
and Default times with given survival probability and their \mathbb{F} -martingale
decomposition formula. *In preparation.*

El Karoui, N., Jeanblanc, M. and Jiao, Y. (2009). What happens after a
default: the conditional density approach. *Forthcoming SPA* 120 (2010) 1011-1032.

Related works:

Brody D.C. and Hughston L.P. (2002): Entropy and information in the interest rate term structure, *Quantitative Finance*, 2, 70-80.

Papapantoleon, A. (2009): Old and new approaches to Libor modeling, preprint.

Filipovic, D., L. Overbeck and T. Schmidt (2009): Dynamic CDO term structure modelling, preprint

Mathematical Model

A filtered probability space $(\Omega, \mathcal{A}, \mathbb{F}, \mathbb{P})$ is given, as well as a random time τ . The default process is $H_t = \mathbb{1}_{\tau \leq t}$, the associated filtration is $\mathbb{H} = (\mathcal{H}_t = \sigma(t \wedge \tau), t \geq 0)$. The filtration \mathbb{G} is defined as $\mathcal{G}_t = \mathcal{F}_t \vee \mathcal{H}_t$. The \mathbb{G} -intensity of τ is the process $\lambda^{\mathbb{G}}$ such that

$$M_t = H_t - \int_0^t \lambda_s^{\mathbb{G}} ds$$

is a \mathbb{G} -martingale. There exists an \mathbb{F} -adapted process $\lambda^{\mathbb{F}}$ such that

$$M_t = H_t - \int_0^{t \wedge \tau} \lambda_s^{\mathbb{F}} ds$$

If $X \in \mathcal{F}_T$, and $G_t = \mathbb{P}(\tau > t | \mathcal{F}_t)$, then

$$\mathbb{E}(X \mathbb{1}_{T < \tau} | \mathcal{G}_t) = \mathbb{1}_{t < \tau} \frac{1}{G_t} \mathbb{E}(X G_T | \mathcal{F}_t)$$

One can think that the knowledge of λ and G will allow us to have the knowledge of the conditional law of τ . We shall show that this is not the case.

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Intensity models

Models with a given intensity are constructed as follows.

Let λ be a given \mathbb{F} -adapted positive process and Θ a random variable independent of \mathcal{F}_∞ , with unit exponential law. Then

$$\tau = \inf\{t : \int_0^t \lambda_s ds \geq \Theta\}$$

has intensity equal to λ .

In that model, $\mathbb{P}(\tau > u | \mathcal{F}_t) = \mathbb{E}(e^{-\Lambda_u} | \mathcal{F}_t)$ and immersion property holds:

$$\begin{aligned} \mathbb{P}(\tau > t | \mathcal{F}_t) &= \mathbb{P}(\tau > t | \mathcal{F}_\infty) = e^{-\Lambda_t} \\ \mathbb{E}(X | \mathcal{F}_t) &= \mathbb{E}(X | \mathcal{G}_t), \forall X \in \mathcal{F}_\infty \end{aligned}$$

Under immersion property, one has

$$p_t(u) du := \mathbb{P}(\tau \in du | \mathcal{F}_t) = \mathbb{E}(\lambda_u e^{-\Lambda_u} | \mathcal{F}_t) du$$

and we note that $p_t(u) = p_u(u), \forall t \geq u$.

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We now construct probabilities \mathbb{Q} equivalent to \mathbb{P} such that τ has intensity λ , and where immersion does not hold, hence, for $t > u$, the density $p_t^{\mathbb{Q}}(u)$ is not determined in terms of the intensity.

Let $p_t(u)du = \mathbb{P}(\tau \in du | \mathcal{F}_t)$ and $z(u)$ a family of processes such that

(i) $(z_t(u), t \geq u)$ are positive \mathbb{F} -martingales.

Define, for z positive \mathbb{F} -adapted process

$$Z_t^{\mathbb{G}} = z_t \mathbb{1}_{\{\tau > t\}} + z_t(\tau) \mathbb{1}_{\{\tau \leq t\}}$$

and let

$$Z_t^{\mathbb{F}} := \mathbb{E}(Z_t^{\mathbb{G}} | \mathcal{F}_t) = z_t G_t + \int_0^t z_t(u) p_t(u) du$$

be its \mathbb{F} -projection. Assume that

(ii) $Z^{\mathbb{F}}$ is a \mathbb{F} -martingale.

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Proof: (we assume here that G is continuous.) Let $s < t$.

$$\mathbb{E}(Z_t^{\mathbb{G}}|\mathcal{G}_s) = \mathbb{E}(z_t \mathbb{1}_{\tau > t} | \mathcal{G}_s) + \mathbb{E}(z_t(\tau) \mathbb{1}_{s < \tau \leq t} | \mathcal{G}_s) + \mathbb{E}(z_t(\tau) \mathbb{1}_{\tau \leq s} | \mathcal{G}_s) = I_1 + I_2.$$

For I_1 , we apply the standard formula

$$I_1 = \mathbb{1}_{\tau > s} \frac{1}{G_s} \mathbb{E}(Z_t G_t | \mathcal{F}_s) + \mathbb{1}_{\tau > s} \frac{1}{G_s} \mathbb{E}(z_t(\tau) \mathbb{1}_{s < \tau \leq t} | \mathcal{F}_s),$$

For I_2 , we obtain

$$I_2 = \mathbb{E}(z_t(\tau) \mathbb{1}_{\tau \leq s} | \mathcal{G}_s) = \mathbb{1}_{\tau \leq s} \mathbb{E}(z_t(u) | \mathcal{F}_s)_{u=\tau} = \mathbb{1}_{\tau \leq s} (z_s(u))_{u=\tau} = \mathbb{1}_{\tau \leq s} z_s(\tau),$$

where the first equality holds under the immersion hypothesis and the second follows from (i). It thus suffices to show that $I_1 = Z_s \mathbb{1}_{\tau > s}$.

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Condition (ii) yields

$$\mathbb{E}(z_t G_t | \mathcal{F}_s) + \mathbb{E}(z_t(\tau) \mathbb{1}_{\tau \leq t} | \mathcal{F}_s) - \mathbb{E}(z_s(\tau) \mathbb{1}_{\tau \leq s} | \mathcal{F}_s) = z_s G_s.$$

Therefore,

$$I_1 = \mathbb{1}_{\tau > s} \frac{1}{G_s} \left(z_s G_s + \mathbb{E}((z_s(\tau) - z_t(\tau)) \mathbb{1}_{\tau \leq s} | \mathcal{F}_s) \right) = z_s \mathbb{1}_{\tau > s},$$

where the last equality holds since

$$\mathbb{E}((z_s(\tau) - z_t(\tau)) \mathbb{1}_{\tau \leq s} | \mathcal{F}_s) = \mathbb{1}_{\tau \leq s} \mathbb{E}((z_s(u) - z_t(u)) | \mathcal{F}_s)_{u=\tau} = 0.$$

For the last equality in the formula above, we have again used condition (i).

We assume (w.l.g.) that $Z_0^{\mathbb{G}} = 1$.

Let \mathbb{Q} be the probability measure defined on \mathcal{G}_t by $d\mathbb{Q} = Z_t^{\mathbb{G}} d\mathbb{P}$.

We assume that $z_t(t) = z_t$ (so that the RN density has no jump at time τ).

Then, for $t \geq \theta$,

$$p_t^{\mathbb{Q}}(\theta) = p_t(\theta) \frac{z_t(\theta)}{Z_t^{\mathbb{F}}},$$

and the \mathbb{Q} -conditional survival process is defined by

$$\mathbb{Q}(\tau > t | \mathcal{F}_t) = e^{-\Lambda_t} \frac{z_t}{Z_t^{\mathbb{F}}} = N_t^{\mathbb{Q}} e^{-\Lambda_t}$$

(in particular, the \mathbb{Q} -intensity and the \mathbb{P} -intensity are the same).

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Proof: For $t > u$,

$$\mathbb{Q}(\tau > u | \mathcal{F}_t) = \frac{1}{\mathbb{E}(Z_t^{\mathbb{G}} | \mathcal{F}_t)} \mathbb{E}(Z_t^{\mathbb{G}} \mathbb{1}_{u < \tau} | \mathcal{F}_t)$$

$$\begin{aligned} \mathbb{E}(Z_t^{\mathbb{G}} \mathbb{1}_{u < \tau} | \mathcal{F}_t) &= \mathbb{E}(Z_t^{\mathbb{G}} \mathbb{1}_{t < \tau} | \mathcal{F}_t) + \mathbb{E}(Z_t^{\mathbb{G}} \mathbb{1}_{u < \tau \leq t} | \mathcal{F}_t) = z_t G_t + \mathbb{E}(z_t(\tau) \mathbb{1}_{u < \tau \leq t} | \mathcal{F}_t) \\ &= z_t G_t + \int_u^t z_t(v) p_t(v) dv \end{aligned}$$

and the density follows by differentiation. The form of the intensity ($\lambda_t^{\mathbb{G}} = \frac{p_t^{\mathbb{G}}(t)}{G_t^{\mathbb{G}}}$) follows. Indeed, if $G_t = \mu_t - A_t$ is the Doob-Meyer decomposition of G , $A_t = \int_0^t p_u(u) du$ and the intensity is $\lambda_t dt = \frac{dA_t}{G_t}$.

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Construction of a random time with given conditional law

Let $p(u)$ a family of positive \mathbb{F} -martingales such that

$$\int_0^\infty p_t(u) du = 1, \forall t$$

One can construct (on an extended space) a probability \mathbb{Q} and a random time τ such that

$$\begin{aligned}\mathbb{Q}|_{\mathcal{F}_t} &= \mathbb{P}|_{\mathcal{F}_t} \\ \mathbb{Q}(\tau \in du | \mathcal{F}_t) &= p_t(u) du\end{aligned}$$

as follows:

- Construct \mathbb{Q}^* and τ such that τ is independent from \mathcal{F}_∞ and $\mathbb{Q}(\tau \in du) = p_0(u) du$
- Set $d\mathbb{Q}|_{\mathcal{F}_t \vee \sigma(\tau)} = (p_t(\tau))^{-1} d\mathbb{Q}^*$

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Construction of a random time with given Conditional Survival Probability

Construct (on an extended space) a probability \mathbb{Q} and a random time τ such that

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where G is a given \mathbb{F} -supermartingale. One recall that any supermartingale admits a multiplicative decomposition as $G_t = N_t D_t = N_t e^{-\Lambda_t}$ where D (resp. Λ) is decreasing (resp. increasing) In what follows, we assume that G is continuous, and $0 \leq G_t < 1$ for $t > 0$.

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Construction of a random time with given Conditional Survival Probability

Let us start with a model in which $\mathbb{P}(\tau > t | \mathcal{F}_t) = e^{-\Lambda_t}$, where $\Lambda_t = \int_0^t \lambda_s ds$ and let N be an \mathbb{F} -local martingale such that $0 \leq N_t e^{-\Lambda_t} \leq 1$.

There exists a \mathbb{G} -martingale L such that, setting $d\mathbb{Q} = L d\mathbb{P}$

- (i) $\mathbb{Q}|_{\mathcal{F}_\infty} = \mathbb{P}|_{\mathcal{F}_\infty}$
- (ii) $\mathbb{Q}(\tau > t | \mathcal{F}_t) = N_t e^{-\Lambda_t}$

The \mathbb{G} -adapted process L

$$L_t = \ell_t \mathbb{1}_{t < \tau} + \ell_t(\tau) \mathbb{1}_{\tau \leq t}$$

is a martingale if for any u , $(\ell_t(u), t \geq u)$ is a martingale and if $\mathbb{E}(L_t | \mathcal{F}_t)$ is a \mathbb{F} -martingale. Then, (i) is satisfied if

$$1 = \mathbb{E}(L_t | \mathcal{F}_t) = \ell_t G_t + \int_0^t \ell_t(u) \lambda_u e^{-\Lambda_u} du$$

and (ii) implies that $\ell = N$ and $\ell_t(t) = \ell_t$.

Conditional Survival Probability

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Conditional Survival Probability

Let us start with a model in which $\mathbb{P}(\tau > t | \mathcal{F}_t) = e^{-\Lambda_t}$, where $\Lambda_t = \int_0^t \lambda_s ds$ and let N be an \mathbb{F} -local martingale such that $0 \leq N_t e^{-\Lambda_t} \leq 1$.

There exists a \mathbb{G} -martingale L such that, setting $d\mathbb{Q} = L d\mathbb{P}$

$$(i) \quad \mathbb{Q}|_{\mathcal{F}_\infty} = \mathbb{P}|_{\mathcal{F}_\infty}$$

$$(ii) \quad \mathbb{Q}(\tau > t | \mathcal{F}_t) = N_t e^{-\Lambda_t}$$

The \mathbb{G} -adapted process L

$$L_t = \ell_t \mathbb{1}_{t < \tau} + \ell_t(\tau) \mathbb{1}_{\tau \leq t}$$

is a martingale if for any u , $(\ell_t(u), t \geq u)$ is a martingale and if $\mathbb{E}(L_t | \mathcal{F}_t)$ is a \mathbb{F} -martingale. Then, (i) is satisfied if

$$1 = \mathbb{E}(L_t | \mathcal{F}_t) = \ell_t e^{-\Lambda_t} + \int_0^t \ell_t(u) \lambda_u e^{-\Lambda_u} du$$

and (ii) implies that $\ell = N$ and $\ell_t(t) = \ell_t$.

It remains to find a family of martingales $\ell(u)$ such that

$$\begin{aligned}\ell_u(u) &= N_u \\ 1 &= N_t e^{-\Lambda_t} + \int_0^t \ell_t(u) \lambda_u e^{-\Lambda_u} du\end{aligned}$$

We choose

$$\ell_t(u) = \frac{N_u}{1 - G_u} (1 - G_t) \exp \left(- \int_u^t \frac{G_s}{1 - G_s} \lambda_s ds \right)$$

Then, $\mathbb{Q}[\tau \leq u | \mathcal{F}_t] = M_t^u$ for $0 \leq u \leq t \leq \infty$ where

$$M_t^u = (1 - G_t) \exp \left(- \int_u^t \frac{G_s}{1 - G_s} \lambda_s ds \right) \quad 0 \leq u \leq t \leq \infty,$$

One can also construct other martingales M^u which give a solution (i.e., families of $[0, 1]$ -valued martingales such that $u \rightarrow M_t^u$ is decreasing and $M_t^t = 1 - G_t$).

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Cox processes

Let λ be a strictly positive \mathbb{F} -adapted process, and $\Lambda_t = \int_0^t \lambda_s ds$.

Let Θ be a strictly positive random variable **whose conditional distribution w.r.t. \mathbb{F} admits a density** w.r.t. the Lebesgue measure, i.e., there exists a family of $\mathcal{F}_t \otimes \mathcal{B}(\mathbb{R}_+)$ -measurable functions $\gamma_t(u)$ such that $\mathbb{P}(\Theta > \theta | \mathcal{F}_t) = \int_\theta^\infty \gamma_t(u) du$.

Let $\tau = \inf\{t > 0 : \Lambda_t \geq \Theta\}$.

Then τ admits the density

$$p_t(\theta) = \lambda_\theta \gamma_t(\Lambda_\theta) \text{ if } t \geq \theta \quad \text{and} \quad p_t(\theta) = \mathbb{E}[\lambda_\theta \gamma_\theta(\Lambda_\theta) | \mathcal{F}_t] \text{ if } t < \theta.$$

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Proof: By definition and by the fact that Λ is strictly increasing and absolutely continuous, we have for $t \geq \theta$,

$$\mathbb{P}(\tau > \theta | \mathcal{F}_t) = \mathbb{P}(\Theta > \Lambda_\theta | \mathcal{F}_t) = \int_{\Lambda_\theta}^{\infty} \gamma_t(u) du = \int_{\theta}^{\infty} \gamma_t(\Lambda_u) d\Lambda_u = \int_{\theta}^{\infty} \gamma_t(\Lambda_u) \lambda_u du,$$

which implies $p_t(\theta) = \lambda_\theta \gamma_t(\Lambda_\theta)$. The martingale property of p gives the whole density.

Conversely, if we are given a density p , and hence an associated process $\Lambda_t = \int_0^t \lambda_s ds$ with $\lambda_s = \frac{p_s(s)}{G_s}$, then it is possible to find a threshold Θ such that τ has p as density.

We denote by Λ^{-1} the inverse of the strictly increasing process λ .

We let $\Lambda_t = \int_0^t \frac{p_s(s)}{G_s} ds$ and $\Theta = \Lambda_\tau$. Then Θ has a density γ with respect to \mathbb{F} given by

$$\gamma_t(\theta) = \mathbb{E} \left[p_{t \vee \Lambda_\theta^{-1}}(\Lambda_\theta^{-1}) \frac{1}{\lambda_{\Lambda_\theta^{-1}}} | \mathcal{F}_t \right].$$

Proof: We set $\Theta = \Lambda_\tau$ and compute the density of Θ w.r.t. \mathbb{F}

$$\begin{aligned}
\mathbb{P}(\Theta > \theta | \mathcal{F}_t) &= \mathbb{P}(\Lambda_\tau > \theta | \mathcal{F}_t) = \mathbb{P}(\tau > t, \Lambda_\tau > \theta | \mathcal{F}_t) + \mathbb{P}(\tau \leq t, \Lambda_\tau > \theta | \mathcal{F}_t) \\
&= \mathbb{E}\left[-\int_t^\infty \mathbb{1}_{\{\Lambda_u > \theta\}} dG_u | \mathcal{F}_t\right] + \int_0^t \mathbb{1}_{\{\Lambda_u > \theta\}} p_t(u) du \\
&= \mathbb{E}\left[\int_t^\infty \mathbb{1}_{\{\Lambda_u > \theta\}} p_u(u) du | \mathcal{F}_t\right] + \int_0^t \mathbb{1}_{\{\Lambda_u > \theta\}} p_t(u) du
\end{aligned}$$

where the last equality comes from the fact that $(G_t + \int_0^t p_u(u) du, t \geq 0)$ is an \mathbb{F} -martingale. Note that since the process Λ is continuous and strictly increasing, also is its inverse. Hence

$$\begin{aligned}
&\mathbb{E}\left[\int_\theta^\infty p_{\Lambda_s^{-1} \vee t}(\Lambda_s^{-1}) \frac{1}{\lambda_{\Lambda_s^{-1}}} ds | \mathcal{F}_t\right] = \mathbb{E}\left[\int_{\Lambda_\theta^{-1}}^\infty p_{s \vee t}(s) \frac{1}{\lambda_s} d\Lambda_s | \mathcal{F}_t\right] \\
&= \mathbb{E}\left[\int_0^\infty \mathbb{1}_{\{s > \Lambda_\theta^{-1}\}} p_{s \vee t}(s) ds | \mathcal{F}_t\right] = \mathbb{E}\left[\int_0^\infty \mathbb{1}_{\{\Lambda_s > \theta\}} p_{s \vee t}(s) ds | \mathcal{F}_t\right],
\end{aligned}$$

which equals $\mathbb{P}(\Theta > \theta | \mathcal{F}_t)$.

Defaultable Zero-Coupon Bonds

A defaultable zero-coupon with maturity T associated with the default time τ is an asset which pays one monetary unit at time T if (and only if) the default has not occurred before T . We assume that \mathbb{P} is the pricing measure.

$$D(t, T) := \mathbb{P}(\tau > T \mid \mathcal{G}_t) = \mathbb{1}_{\{\tau > t\}} \frac{\mathbb{P}(\tau > T \mid \mathcal{F}_t)}{G_t} = \mathbb{1}_{\{\tau > t\}} \frac{\mathbb{E}_{\mathbb{P}}(N_T e^{-\Lambda_T} \mid \mathcal{F}_t)}{G_t}$$

However, using a change of probability, one can get rid of the martingale part N , assuming that there exists p such that

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Let \mathbb{P}^* be defined as

$$d\mathbb{P}^*|_{\mathcal{G}_t} = Z_t^* d\mathbb{P}|_{\mathcal{G}_t}$$

where Z^* is the (\mathbb{P}, \mathbb{G}) -martingale defined as

$$Z_t^* = \mathbb{1}_{\{t < \tau\}} + \mathbb{1}_{\{t \geq \tau\}} \lambda_\tau e^{-\Lambda_\tau} \frac{N_t}{p_t(\tau)}$$

Then,

- (a) Immersion property holds under \mathbb{P}^* ,
- (b) $d\mathbb{P}^*|_{\mathcal{F}_t} = N_t d\mathbb{P}|_{\mathcal{F}_t}$
- (c) \mathbb{P}^* and \mathbb{P} coincide on \mathcal{G}_τ .

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Proof: We prove first that $d\mathbb{P}^*|_{\mathcal{F}_t} = N_t d\mathbb{P}|_{\mathcal{F}_t}$

$$\mathbb{E}_{\mathbb{P}}(Z_t^*|\mathcal{F}_t) = G_t + \int_0^t \lambda_u e^{-\Lambda_u} \frac{N_t}{p_t(u)} p_t(u) du = N_t e^{-\Lambda_t} + N_t(1 - e^{-\Lambda_t}) = N_t$$

We compute the \mathbb{P}^* conditional law of τ . For $t > \theta$,

$$\begin{aligned} \mathbb{P}^*(\theta < \tau|\mathcal{F}_t) &= \frac{1}{N_t} \mathbb{E}_{\mathbb{P}}(Z_t^* \mathbb{1}_{\{\theta < \tau\}}|\mathcal{F}_t) = \frac{1}{N_t} \mathbb{E}_{\mathbb{P}}(\mathbb{1}_{\{t < \tau\}} + \mathbb{1}_{\{t \geq \tau > \theta\}} \lambda_{\tau} e^{-\Lambda_{\tau}} \frac{N_t}{p_t(\tau)}|\mathcal{F}_t) \\ &= \frac{1}{N_t} \left(N_t e^{-\Lambda_t} + \int_{\theta}^t \lambda_u e^{-\Lambda_u} \frac{N_t}{p_t(u)} p_t(u) du \right) = e^{-\Lambda_{\theta}} \end{aligned}$$

which proves that immersion holds true under \mathbb{P}^* , and the intensity of τ is the same under \mathbb{P} and \mathbb{P}^* . It follows that

$$\mathbb{E}^*(X \mathbb{1}_{\{T < \tau\}}|\mathcal{G}_t) = \mathbb{1}_{\{t < \tau\}} \frac{1}{e^{-\Lambda_t}} \mathbb{E}^*(e^{-\Lambda_{\tau}} X|\mathcal{F}_t) = \mathbb{E}_{\mathbb{P}}(X \mathbb{1}_{\{T < \tau\}}|\mathcal{G}_t)$$

Note that, if the intensity is the same under \mathbb{P} and \mathbb{P}^* , its dynamics under \mathbb{P}^* will involve a change of driving process, since \mathbb{P} and \mathbb{P}^* do not coincide on \mathcal{F}_{∞} .

Let us now study the pricing of a recovery. Let Z be an \mathbb{F} -predictable bounded process.

$$\begin{aligned}\mathbb{E}_{\mathbb{P}}(Z_{\tau} \mathbb{1}_{\{t < \tau \leq T\}} | \mathcal{G}_t) &= \mathbb{1}_{\{t < \tau\}} \frac{1}{G_t} \mathbb{E}_{\mathbb{P}}\left(-\int_t^T Z_u dG_u | \mathcal{F}_t\right) \\ &= \mathbb{1}_{\{t < \tau\}} \frac{1}{G_t} \mathbb{E}_{\mathbb{P}}\left(\int_t^T Z_u N_u \lambda_u e^{-\Lambda_u} du | \mathcal{F}_t\right) \\ &= \mathbb{1}_{\{t < \tau\}} \frac{1}{e^{-\Lambda_t}} \mathbb{E}^*\left(\int_t^T Z_u \lambda_u e^{-\Lambda_u} du | \mathcal{F}_t\right) \\ &= \mathbb{E}^*(Z_{\tau} \mathbb{1}_{\{t < \tau \leq T\}} | \mathcal{G}_t)\end{aligned}$$

The problem is different for pricing a recovery paid at maturity, i.e. for $X \in \mathcal{F}_T$

$$\begin{aligned}\mathbb{E}_{\mathbb{P}}(X \mathbb{1}_{\tau < T} | \mathcal{G}_t) &= \mathbb{E}_{\mathbb{P}}(X | \mathcal{G}_t) - \mathbb{E}_{\mathbb{P}}(X \mathbb{1}_{\tau > T} | \mathcal{G}_t) \\ &= \mathbb{E}_{\mathbb{P}}(X | \mathcal{G}_t) - \mathbb{1}_{\{\tau > t\}} \frac{1}{e^{-\Lambda_t}} \mathbb{E}^*(X e^{-\Lambda_T} | \mathcal{F}_t)\end{aligned}$$

If both quantities $\mathbb{E}_{\mathbb{P}}(X \mathbb{1}_{\tau < T} | \mathcal{G}_t)$ and $\mathbb{E}^*(X \mathbb{1}_{\tau < T} | \mathcal{G}_t)$ are the same, this would imply that $\mathbb{E}_{\mathbb{P}}(X | \mathcal{G}_t) = \mathbb{E}^*(X | \mathcal{F}_t)$ i.e., immersion holds under \mathbb{P} .

Misspecification of the Information

In this section, we point out that the price of a derivative product written on a default τ depends strongly on the other default and the hedging strategies have to be constructed with the full observation. Let us study the following toy model

Two default times τ_1, τ_2 let $G(t, s) = \mathbb{P}(\tau_1 > t, \tau_2 > s)$

We consider two filtrations \mathbb{H}^1 and $\mathbb{H} = \mathbb{H}^1 \vee \mathbb{H}^2$

The price of a DZC is

$$\begin{aligned}\bar{D}^1(t, T) &= \mathbb{P}(\tau_1 > T | \mathcal{H}_t^1) = (1 - H_t^1) \frac{G(T, 0)}{G(t, 0)} \\ D^1(t, T) &= \mathbb{P}(\tau_1 > T | \mathcal{H}_t) = (1 - H_t^1) \left((1 - H_t^2) \frac{G(T, t)}{G(t, t)} + H_t^2 \frac{\partial_2 G(T, \tau_2)}{\partial_2 G(t, \tau_2)} \right)\end{aligned}$$

The un-informed agent knows only \mathbb{H}^1 . He will hedge the contingent claim $C = h(\tau_1)\mathbb{1}_{\tau_1 > T} + k\mathbb{1}_{\tau_1 > T}$ thinking the market is complete, with an initial wealth $x = \mathbb{E}(C)$ buying ζ_t DZC, so that

$$\hat{X}_T := x + \int_0^T \zeta_s d\hat{D}(t, T) = C$$

and he will invest $\zeta_t^0 = X_t - \zeta_t \hat{D}(t, T)$ in the savings account in a self financing way. However, his "real" wealth will be $X_t = \zeta_t^0 + \zeta_t D(t, T)$ and the strategy is not self-financing. The cost of refinancing is

$$dX_t - \zeta_t dD(t, T) = d\zeta_t^0 + D(t, T)d\zeta_t^0 + d\langle \zeta, D(\cdot, T) \rangle_t$$

If he uses a self financing strategy, his terminal wealth will be

$X_T^* = x + \int_0^T \zeta_t dD(t, T)$ and the associated cost is

$C - X_T^* = \int_0^T \zeta_t (d\hat{D}(t, T) - dD(t, T))$. One has $\mathbb{E}(C - X_T^*) = 0$

Thank you for your attention