

Backward SDEs, Lecture II

Existence, Stability, and Numerical Methods

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- 1 Existence Results
- 2 Reflected BSDEs
- 3 Numerical methods
- 4 Computations of the conditional expectations

Backward Stochastic Differential Equation

- ▶ Standard filtered probability space $(\Omega, \mathcal{F}, (\mathcal{F}_t), 0 \leq t \leq T, \mathbb{P})$, supporting a **standard BM** $W \in \mathbb{R}^n$.
- ▶ A non anticipating **coefficient** $f(t, \omega, y, z)$ defined on $(\Omega \times \mathbb{R}^+, \mathbb{R}^d \times \mathbb{R}^{d \times n})$, a terminal **condition** $\xi_T \in \mathcal{F}_T$

Definition of BSDE solution

A solution of BSDE(f, ξ_T), is a par of non anticipating processes $(Y, Z) \in \mathbb{R}^d \times \mathbb{R}^{d \times n}$ such that

- ▶ $Y_t = \xi_T + \int_t^T f(s, Y_s, Z_s)ds - \int_t^T Z_s dW_s,$
 - ▶ or equivalently $-dY_t = f(t, Y_t, Z_t)ds - Z_t dW_t, \quad Y_T = \xi_T$
 - ▶ with minimal integrability condition, $\int_0^T (|f(t, Y_t, Z_t)| + |Z_t|^2)dt < \infty$ a.s.
-
- ▶ Existence, Uniqueness? : in which spaces of processes,...
 - ▶ Properties? : Stability, Comparison Theorem.....

Doob Inequalities

Notation for the running maximum : $\max |M|_T = \sup_{0,T} |M|_s$

Continuous Martingale : a priori estimates

► Doob inequalities :

$$\mathbb{E}[\max |M|_T^2] \leq c \mathbb{E}[|M_T|^2] \leq C \mathbb{E}[\max |M|_T^2]$$

Should be read in both directions ($A \leq B \leq C$)

- $B \implies A$ is a Backward inequality
- $C \implies B$ is a Forward inequality

► Burkholder, Davis Gundy inequalities

Let $\langle M \rangle$ be the a quadratic variation of M , then for any $p > 0$

$$\mathbb{E}[\max |M|_T^p] \leq c_p \mathbb{E}[|M_T|^p/2] \leq C_p \mathbb{E}[\max |M|_T^p]$$

Representation Theorem

A priori Forward or Backward Estimates

Weighted \mathbb{H}_T^2 space

- ▶ **Forward** \mathbb{H}_c^2 , defined as \mathbb{H}_T^2 with the semi-norm

$$\|X\|_c^2 = \max(e^{-2ct} \mathbb{E}[\max |X_t|^2])_T$$
- ▶ **Backward** \mathbb{H}_β^2 , defined as \mathbb{H}_T^2 with the semi-norm

$$\|X\|_\beta^2 = \max(e^{2\beta t} \mathbb{E}[\max |X_t|^2])_T$$

Estimates of $F_t^T = \int_t^T f_s ds$ a finite variation process.

- ▶ **Forward**

$$|F_t^T|^2 = \left| \int_t^T e^{sc/2} (e^{-sc/2} f_s) ds \right|^2 \leq e^{cT} \frac{1}{c} \int_t^T e^{-cs} |f_s|^2 ds$$
- ▶ **Backward**

$$|F_t^T|^2 = \left| \int_t^T e^{-s\beta/2} (e^{s\beta/2} f_s) ds \right|^2 \leq e^{-\beta t} \frac{1}{\beta} \int_t^T e^{s\beta} |f_s|^2 ds$$

Semimartingale Estimates

Let $x_T = x_t - \int_t^T f_s ds - \int_t^T \eta_s dW_s$ a Itô's semimartingale

► Forward

Since $x_t = x_0 - \int_0^t f_s ds - \int_0^t \eta_s dW_s$, then

$|x_t| \leq |y_0| + |F_0^t| + \max |\eta \cdot W|_t$. By the Doob inequality,

$$e^{-ct} \mathbb{E}[\max |x_t|^2] \leq \mathbb{E}\left[e^{-ct} |x_0|^2 + \frac{1}{c} \int_0^t e^{-cs} (|f_s|^2 + |\eta_s|^2) ds\right]$$

$$\|x\|_c^2 \leq 2\mathbb{E}\left[e^{-cT} |x_0|^2 + \frac{1}{c} \int_0^T e^{-sc} (|f_s|^2 + |\eta_s|^2) ds\right]$$

► Backward

By Doob inequality, since $|x_t| \leq \mathbb{E}[|x_T| + |F_t| | \mathcal{F}_t]$,

$$e^{t\beta} |x_t| \leq \mathbb{E}\left[\left(e^{T\beta/2} |x_T| + \frac{1}{\beta} \int_t^T e^{s\beta} |f_s|^2 ds\right)^{1/2} | \mathcal{F}_t\right]$$

$$\begin{cases} \|x\|_\beta^2 & \leq 4\mathbb{E}[e^{T\beta} |x_T|^2 + \frac{1}{\beta} \int_0^T e^{s\beta} |f_s|^2 ds] \\ \|\eta \cdot W\|_\beta^2 & \leq K \left[\mathbb{E}[e^{T\beta} |x_T|^2 + \frac{1}{\beta} \int_0^T e^{s\beta} |f_s|^2 ds] \right] \end{cases}$$

Lipschitz Assumptions

Forward Assumptions

- ▶ $F(t, [x]_t)$, and $G(t, [x]_t)$ (path dependency) in \mathbb{L}^2
- ▶ Uniformly Lipschitz i.e, there exists $K > 0$ s.t a.s
$$|F(t, [x^1]_t) - F(t, [x^2]_t)| + |G(t, [x^1]_t) - G(t, [x^2]_t)| \leq K |[x_1 - x_2]|$$

Backward Assumptions

- ▶ Standard data $(f, \xi) : \int_0^T |f(t, 0, 0)|^2 ds, \xi \in \mathbb{L}^2$
- ▶ f is uniformly lipschitz, i.e., there exists $C > 0$ s.t a.s
$$|f(t, y_1, z_1) - f(t, y_2, z_2)| \leq C (|y_1 - y_2| + |z_1 - z_2|)$$

Notations : given two coefficients f^1, f^2 ,

- ▶ $\delta Y_t = Y_t^1 - Y_t^2, \delta Z_t = Z_t^1 - Z_t^2$
- ▶ $\delta_2 f_t = f^1(t, y_2, z_2) - f^2(t, y_2, z_2), \delta_2 F_t = \delta_2 f_t(Y_t^2, Z_t^2)$

Solutions via Picard Approximations

► Forward Lipschitz SDE

$$dX_t = F(t, [X]_t)dt + G(t, [X]_t)dW_t$$

- General filtration
- Standard \mathbb{L}^2 multi-dim data (X_0, F, G) , uniformly Lipschitz.
- Existence and Uniqueness
 - \exists a unique solution in \mathbb{H}_T^2

► Backward Lipschitz SDE

$$-dY_t = f(t, Y_t, Z_t)dt, -Z_t.dW_t, \\ Y_T = \xi_T$$

- Brownian Filtration
- Standard \mathbb{L}^2 multi-dim data (f, ξ_T) , uniformly Lipschitz.
- Existence and Uniqueness
 - \exists a unique pair $(Y, Z) \in \mathbb{H}^2$

In the both cases, the Picard sequence converges uniformly in the right \mathbb{H}_T^2 space to the solution with an exponential speed. The estimates are uniform in the boundary conditions.

General Markovian Setting

Let X be a **diffusion process** on a general filtered probability space, and \mathcal{B}_e be the σ -field on \mathbb{R}^n generated by $\mathbb{E} \int_t^T \phi(s, X_s^{t,x}) ds$ where ϕ is a continuous bounded. Let $(f, \Psi) \in \mathcal{B}_e$ be squared integrable ($\mathbb{E} \int_0^T f^2(s, X_s^{t,x}) ds < +\infty$; $\mathbb{E}[\Psi^2(X_T^{t,x})] < +\infty$),

► **Markovian representation of the solution**[CJPS]

The semimartingale $Y_s^{t,x} = \mathbb{E}[\Psi(X_T^{t,x}) + \int_s^T f(r, X_r^{t,x}) dr | \mathcal{F}_s]$ admits a continuous version given by $u(s, X_s^{t,x})$ with $u(t, x) = Y_t^{t,x} \in \mathcal{B}_e$

► **Markovian representation of the martingale** Moreover,

$u(t, x) + \int_t^s f(r, X_r^{t,x}) dr + Y_s^{t,x} = U_s^{t,x}$ is an additive martingale with the following representation depending on $d(t, x) \in \mathcal{B}_e$,

$$U_s^{t,x} = \int_t^s \underbrace{d(r, X_r^{t,x})^* \sigma(r, X_r^{t,x})}_{Z_r^{t,x}*} dW_r ; t \leq s$$

Markovian BSDEs

Let X be a **diffusion process** and the associated BSDE :

$$-dY_s = f(s, X_s^{t,x}, Y_s, Z_s)ds - Z_s^* dW_s, \quad Y_T = \Psi(X_T^{t,x})$$

- **General setting** : Thanks to Picard approximates, there exists

$u(t, x), d(t, x) \in \mathcal{B}_e$ such that

$$Y_s = u(s, X_s^{t,x}), Z_s = d(s, X_s^{t,x})^* \sigma(s, X_s^{t,x}).$$

- **PDE solution in one dimensional case**

Let \mathcal{L}^X the elliptic operator associated with the diffusion X .

Then, under mild regularity assumptions, u is a viscosity

solution of the **HJB Type PDE**

$$\begin{cases} \partial_t u(t, x) + \mathcal{L}v(t, x) + f(t, x, u(t, x), \partial_x u(t, x)\sigma(t, x)) = 0 \\ u(T, x) = \Psi(x). \end{cases}$$

Then, $d(t, x)$ plays the role of $\partial_x u$ the gradient of u .

proof is provided by the strict comparison theorem.

Linear growth assumption, $d=1$

For simplicity, we assume that $f(t, 0, 0) = 0$

Linear growth : $|f(t, y, z)| \leq g_\mu(y, z) = a|y| + \mu|z|S$ Let \overline{Y}^μ the solution of the Lipschitz BSDE with coefficient g_μ and \underline{Y}^μ the process $-\overline{Y}^\mu(-\xi_T)$. Uniform bounds

Then any square integrable solution (Y, Z) of BSDE(f) with linear growth satisfies

$$\underline{Y}^\mu \leq Y \leq \overline{Y}^\mu$$

Lepeltier, San Martin, '97

There exists a minimal (a maximal)solution to the BSDE with GL continuous coefficient.

General methodology

The different steps of the proof are the following

- ▶ Use a monotone Lipschitz regularisation f^n of f , with **same linear growth**
- ▶ Show that the solutions (Y^n, Z^n) are bounded in L^2 ,

$$\mathbb{E}[\int_0^T |Z_s|^2 ds] \leq C$$
- ▶ Show the control of $\mathbb{E}[\int_0^T |\delta^{i,j} Z_s|^2 ds]$ by $(\mathbb{E}[\int_0^T |\delta^{i,j} Y_s|^2 ds])^{1/2}$
- ▶ Use the monotonicity of the sequence Y^n and the previous estimates to show that Z^n converges strongly in \mathbb{H}^2 to Z , and so Y^n converges uniformly to Y
- ▶ The last step uses the property of the approximating sequence to show that $f^n(t, Y^n, Z^n)$ also converges to $f(t, Y, Z)$

Sketch of the proof

Regularisation by inf convolution

$f^n(x) = \inf_{y \in \mathbb{R}^p} \{f(y) + n|x - y|\}$ is well defined for
 $n \geq \sup(a, \mu) = K$

Key inequality Denote by $Y^{i,j} = \delta^{i,j} Y$ the difference between Y^i and Y^j .

By Itos formula

$$|Y^{i,j}|_t^2 + \int_t^T |Z^{i,j}|_s^2 ds \mathbb{E}_t \left[\int_t^T |Z_s|^2 ds \right]$$

Reflected BSDEs around a regular obstacle

How to maintain a BSDE solution above a given regular obstacle?

Assume $dO_t = U_t dt + V_t dW_t$

. Let (Y, Z) a solution of $\text{BSDE}(f, \xi_T)$

By comparison theorem, if $\xi_T \geq O_T$, and $f(t, O_t, V_t) + U_t \geq 0$, then $Y_t \geq O_t \forall t$

The idea is to push the solution above O_t by adding some "cash", when you need, $f(t, O_t, V_t) + U_t \leq 0$, in a **minimal way**. Working with $Y_t - O_t$, the problem may be rewritten as to push a solution of BSDE above 0.

Definition of Reflected BSDE Above 0

$$\left\{ \begin{array}{l} Y_t = \Phi + \int_t^T f(s, Y_s, Z_s) ds + K_T - K_t - \int_t^T Z_s dW_s, \\ Y_t \geq 0, \\ K \text{ is continuous, increasing, } K_0 = 0 \text{ and } \int_0^T Y_t dK_t = 0. \end{array} \right.$$

The above observation suggests to be looking for a process K absolutely continuous w.r. to $f(t, 0, 0)^- dt$,

$$dK_t = \alpha_t \mathbf{1}_{\{Y_t=0\}} f(t, 0, 0)^- dt, \quad \alpha_t \in [0, 1]$$

Transformation of the problem

The problem is now expressed in terms of α_t .

Regularization

Let ϕ^n a Lipschitz regularization of $\mathbf{1}_{\{y=0\}}$, bounded by 1, and decreasing.

- By the same method that above, one show the same properties holds true, for the BSDEs with $f^n = f + \phi^n(y)$
- to show that the sequence Y^n converges uniformly, and Z^n strongly in L^2 to a pair (Y, Z) , with $Y \geq 0$.
- The only small difficulty is to show that dK_t^n converges to a solution with support $\{Y_t = 0\}$

Applications to optimal stopping problems

General obstacle Lower bound. For any stopping time $\tau \in \mathcal{T}_{t,T}$, one has

$$\begin{aligned} Y_t &= \mathbb{E}\left(Y_t + \int_t^T f(s, Y_s, Z_s) ds + K_\tau - K_t - \int_t^T Z_s dW_s \middle| \mathcal{F}_t\right) \\ &\geq \mathbb{E}(O_t \mathbf{1}_{\tau < T} + \Phi \mathbf{1}_{\tau = T} + \int_t^\tau f(s, Y_s, Z_s) ds \middle| \mathcal{F}_t), \end{aligned}$$

which implies

$$Y_t \geq \operatorname{ess\,sup}_{\tau \in \mathcal{T}_{t,T}} \mathbb{E}(O_\tau \mathbf{1}_{\tau < T} + \Phi \mathbf{1}_{\tau = T} + \int_t^\tau f(s, Y_s, Z_s) ds \middle| \mathcal{F}_t).$$

Equality. The equality holds for $\tau^* = \inf\{u \in [t, T] : Y_u = O_u\} \wedge T$.

Numerical Point of view

New interest for these kind with the swing options, the real options.

⇒ The regular obstacle method is very interesting for numerical methods since

- ▶ it gives an upper approximation (the penalisation app. gives a lower bound).
- ▶ the bounds on the approximated driver depends less on n than for the penalisation scheme.
- ▶ No available estimates on the rate of convergence w.r.t. n .

- ▶ Thanks to Emmanuel Gobet to allows me to use its beautiful presentation of the numerical aspect of BSDEs
 - ▶ The complete presentation may be find on the following site :
 - ▶ [http ://www.cmap.polytechnique.fr?euroschoolmathfi09](http://www.cmap.polytechnique.fr?euroschoolmathfi09)
- Then, go to minicours
- Find the slides of E.Gobet and J.Ma on BSDEs

Our aim :

- ▶ to simulate Y and Z
- ▶ to estimate the error, in order to tune finely the convergence parameters.

Quite intricate and demanding since

- ▶ it is a non-linear problem (the current process dynamics depend on the future evolution of the solution).
- ▶ it involves various deterministic and probabilistic tools (also from statistics).
- ▶ the estimation of the convergence rate is not easy because of the non-linearity, of the loss of independence (mixing of independent simulations).

Strong approximation.

$(X_t^N)_{0 \leq t \leq T}$ is a strong approximation of $(X_t)_{0 \leq t \leq T}$ if

$$\sup_{t \leq T} \|X_t^N - X_t\|_{\mathbb{L}_p} \rightarrow 0 \text{ (or } \|\sup_{t \leq T} |X_t^N - X_t|\|_{\mathbb{L}_p} \rightarrow 0) \text{ as } N \text{ goes to } \infty.$$

Weak approximation. For any test function (smooth or non smooth), one has

$$\mathbb{E}[f(X_T^N)] - \mathbb{E}[f(X_T)] \rightarrow 0 \text{ as } N \text{ goes to } \infty.$$

Examples.

Approximation of SDE : $X_t = x + \int_0^t b(s, X_s)ds + \int_0^t \sigma(s, X_s)dW_s$.

Time discretization using **Euler scheme**. Define $t_k = k\frac{T}{N} = kh$.

$$X_0^N = x, \quad X_{t_{k+1}}^N = X_{t_k}^N + b(t_k, X_{t_k}^N)h + \sigma(t_k, X_{t_k}^N)(W_{t_{k+1}} - W_{t_k}).$$

The simplest scheme to use. Converges at rate $\frac{1}{2}$ for strong approximation and 1 for weak approximation.

Milshtein scheme (not available for arbitrary σ) : rate 1 for both strong and weak approximations.

The BSDE case

We focus mainly on Markovian BSDE :

$Y_t = \Phi(X_T) + \int_t^T f(s, X_s, Y_s, Z_s)ds - \int_t^T Z_s dW_s$, where X is a forward SDE. We know that $Y_t = u(t, X_t)$ and

$Z_t = \nabla_x u(t, X_t) \sigma(t, X_t)$, where u solves a semi-linear PDE

\implies to approximate Y, Z , we need to approximate the function $u(\cdot)$, the gradient of u and the process X

- ▶ $Y_t^N = u^N(t, X_t^N)$,
- ▶ in practice, X^N is always random,
- ▶ although u is deterministic, u^N may be random (e.g. Monte Carlo approximations) : **the randomness may come from two different objects.**

Formal error analysis

$$\begin{aligned}\mathbb{E}|Y_t^N - Y_t| &\leq \mathbb{E}|u^N(t, X_t^N) - u(t, X_t^N)| + \mathbb{E}|u(t, X_t^N) - u(t, X_t)| \\ &\leq \|u^N(t, \cdot) - u(t, \cdot)\|_{\mathbb{L}_\infty} + \|\nabla u\|_{\mathbb{L}_\infty} \mathbb{E}|X_t^N - X_t|.\end{aligned}$$

Two source of error :

- ▶ **strong error** related to $\mathbb{E}|X_t^N - X_t|$.

For the Euler scheme $\mathbb{E}|X_t^N - X_t| = O(N^{-1/2})$.

- ▶ **weak error** related to $\|u^N(t, \cdot) - u(t, \cdot)\|_{\mathbb{L}_\infty}$.

Indeed, to see that this is a weak-type error, take $f \equiv 0$,

$u(t, x) = \mathbb{E}[\Phi(X_T) | X_t = x]$, and the Euler scheme to

approximate the conditional law of X_T : from [BT96], one knows

The grid

Time grid :

$$\pi = \{0 = t_0 < \dots < t_i < \dots < t_N = T\}$$

with non uniform time step : $|\pi| = \max_i(t_{i+1} - t_i)$.

We write $\Delta t_i = t_{i+1} - t_i$ and $\Delta W_{t_i} = W_{t_{i+1}} - W_{t_i}$.

Heuristic derivation

From $Y_{t_i} = Y_{t_{i+1}} + \int_{t_i}^{t_{i+1}} f(s, X_s, Y_s, Z_s) ds - \int_{t_i}^{t_{i+1}} Z_s dW_s$, we derive

$$Y_{t_i} = \mathbb{E}[Y_{t_{i+1}} + \int_{t_i}^{t_{i+1}} f(s, X_s, Y_s, Z_s) ds | \mathcal{F}_{t_i}],$$

$$\mathbb{E}[\int_{t_i}^{t_{i+1}} Z_s ds | \mathcal{F}_{t_i}] = \mathbb{E}[(Y_{t_{i+1}} + \int_{t_i}^{t_{i+1}} f(s, X_s, Y_s, Z_s) ds) \Delta W_{t_i}^* | \mathcal{F}_{t_i}]$$

Discrete backward iteration.

$$\Rightarrow \begin{cases} Z_{t_i}^N = \frac{1}{\Delta t_i} \mathbb{E}[Y_{t_{i+1}}^N - W_{t_i}^* | \mathcal{F}_{t_i}], \\ Y_{t_i}^N = \mathbb{E}[Y_{t_{i+1}}^N + \Delta t_i f(t_i, X_{t_i}^N, Y_{t_{i+1}}^N, Z_{t_{i+1}}^N) | \mathcal{F}_{t_i}] \text{ and } Y_T^N = \psi(X_T^N). \end{cases}$$

The scheme is of **explicit** type.

Implicit scheme

More closely related to the idea of discret BSDE.

$$(\mathbf{Y}_{t_i}^N, \mathbf{Z}_{t_i}^N) = \arg \min_{(Y, Z) \in \mathbb{L}_2(\mathcal{F}_{t_i})} \mathbb{E}[\mathbf{Y}_{t_{i+1}}^N + \Delta t_i f(t_i, \mathbf{X}_{t_i}^N, Y, Z) - Y - Z \Delta \mathbf{W}_{t_i}]^2,$$

with $Y_{t_N}^N = \Phi(X_{t_N}^N)$.

$$\rightarrow \begin{cases} \mathbf{Z}_{t_i}^N = \frac{1}{\Delta t_i} \mathbb{E}[\mathbf{Y}_{t_{i+1}}^N \Delta \mathbf{W}_{t_i}^* | \mathcal{F}_{t_i}], \\ \mathbf{Y}_{t_i}^N = \mathbb{E}[\mathbf{Y}_{t_{i+1}}^N | \mathcal{F}_{t_i}] + \Delta t_i f(t_i, \mathbf{X}_{t_i}^N, \mathbf{Y}_{t_i}^N, \mathbf{Z}_{t_i}^N). \end{cases}$$

Needs a Picard iteration procedure to compute $Y_{t_i}^N$.

Well defined for $|\pi|$ small enough (f Lipschitz).

Define the measure of the squared error

$$\mathcal{E}(Y^N - Y, Z^N - Z) = \max_{0 \leq i \leq N} \mathbb{E}|Y_{t_i}^N - Y_{t_i}|^2 + \sum_{i=0}^{N-1} \int_{t_i}^{t_{i+1}} \mathbb{E}|Z_{t_i}^N - Z_t|^2 dt.$$

Theorem. For a Lipschitz driver w.r.t. (x, y, z) and $\frac{1}{2}$ -Holder w.r.t. t , one has

$$\begin{aligned} \mathcal{E}(Y^N - Y, Z^N - Z) &\leq C(\mathbb{E}|\Phi(X_T^N) - \Phi(X_T)|^2 + \sup_{i \leq N} \mathbb{E}|X_{t_i}^N - X_{t_i}|^2 \\ &\quad + |\pi| + \sum_{i=0}^{N-1} \int_{t_i}^{t_{i+1}} \mathbb{E}|Z_t - \bar{Z}_{t_i}|^2 dt). \end{aligned}$$

where $\bar{Z}_{t_i} = \frac{1}{\Delta t_i} \mathbb{E}(\int_{t_i}^{t_{i+1}} Z_s ds | \mathcal{F}_{t_i})$

Error Analysis

→ Different error contributions :

- ▶ Strong approximation of the forward SDE (depends on the forward scheme and not on the BSDE-problem)
- ▶ Strong approximation of the terminal conditions (depends on the forward scheme and on the BSDE-data Φ)
- ▶ L^2 -regularity of Z (intrinsic to the BSDE-problem).

Diffusion approximation

The easy part : using the Euler scheme

- ▶ $\sup_{i \leq N} |X_{t_i}^N - X_{t_i}|_{\mathbb{L}_2} = O(N^{-1/2})$.
- ▶ If σ does not depend on x , rate $O(N^{-1})$.
- ▶ Otherwise, Milshtein scheme to get N^{-1} -rate.

Strong approximation of the terminal condition

- ▶ If Φ Lipschitz, then $\mathbb{E}|\Phi(X_T^N) - \Phi(X_T)|^2 \leq L_\Phi^2 \mathbb{E}|X_T^N - X_T|^2$.
- ▶ New result if Φ is irregular, using the [approximation theory](#)
Some results of Avikainen [Avi09] for discontinuous function $\Phi(x) = \mathbf{1}_{x \leq a}$.
- ▶ Possible generalization to functions with bounded variation
[\[Avikainen '09\]](#)
- ▶ For intermediate regularity functions, open questions.

$$\mathcal{E}^Z(\pi) = \sum_{i=0}^{N-1} \int_{t_i}^{t_{i+1}} \mathbb{E} |Z_{t_i}^N - Z_t|^2 dt. \text{ Theorem. [Convergence to 0]}$$

Theorem. [Ma, Zhang '02 '04]

Assume a Lipschitz driver f and a Lipschitz terminal condition Φ .

Then Z is a continuous process and $\mathcal{E}^Z(\pi) = O(|\pi|)$ for any time-grid π .

No ellipticity assumption.

Key fact : Z can be represented via a linear BSDE!! It is proved using the Malliavin calculus representation of Z component.

The basics of Malliavin calculus :

Sensitivity of Wiener functionals w.r.t. the BM

For $\xi = \xi(W_t : t \geq 0)$, its Malliavin derivative $(\mathcal{D}_t \xi)_{t \geq 0} \in \mathbb{L}_2(\mathbb{R}^+ \times \Omega, dt \otimes d\mathbb{P})$ is defined as

$$" \mathcal{D}_t \xi = \partial_{dW_t} \xi(W_t : t \geq 0). "$$

Basic rules.

- ▶ If $\xi = \int_0^T h_t dW_t$ with $h \in \mathbb{L}_2(\mathbb{R}^+)$, $\mathcal{D}_t \xi = h_t \mathbf{1}_{t \leq T}$.
- ▶ For smooth random variables $X = g(\int_0^T h_t^1 dW_t, \dots, \int_0^T h_t^n dW_t)$,

$$\mathcal{D}_t X = \sum_{i=1}^n \partial_i g(\dots) h_t^i \mathbf{1}_{t \leq T}.$$
- ▶ Duality relation with adjoint operator \mathcal{D}^* :

Malliavin derivatives of (Y, Z) for smooth data

Theorem.

If $Y_t = \Phi(X_T) + \int_t^T f(s, X_s, Y_s, Z_s)ds - \int_t^T Z_s dW_s$, then for $\theta \leq t \leq T$

$$\begin{aligned} \mathcal{D}_\theta Y_t &= \Phi'(X_T) \mathcal{D}_\theta X_T + \int_t^T [f'_x(s, X_s, Y_s, Z_s) \mathcal{D}_\theta X_s \\ &\quad + f'_y(s, X_s, Y_s, Z_s) \mathcal{D}_\theta Y_s + f'_z(s, X_s, Y_s, Z_s) \mathcal{D}_\theta Z_s] ds - \int_t^T \mathcal{D}_\theta Z_s dW_s \end{aligned}$$

$\implies (\mathcal{D}_\theta Y_t, \mathcal{D}_\theta Z_t)_{t \in [0, T]}$ solves a linear BSDE (for fixed θ).

In addition :

- ▶ Viewing the BSDE as FSDE, one has $Z_t = \mathcal{D}_t Y_t$.
- ▶ Due to $\mathcal{D}_\theta X_t = \nabla X_t [\nabla X_\theta]^{-1} \sigma(\theta, X_\theta)$, we get

$$(\mathcal{D}_\theta Y_t, \mathcal{D}_\theta Z_t) = (\nabla Y_t [\nabla X_\theta]^{-1} \sigma(\theta, X_\theta), \nabla Z_t [\nabla X_\theta]^{-1} \sigma(\theta, X_\theta)),$$

where

$$\begin{aligned} \nabla Y_t = & \Phi'(X_T) \nabla X_T + \int_t^T [f'_X(s, X_s, Y_s, Z_s) \nabla X_s \\ & + f'_Y(s, X_s, Y_s, Z_s) \nabla Y_s + f'_Z(s, X_s, Y_s, Z_s) \nabla Z_s] ds - \int_t^T \nabla Z_s dW_s. \end{aligned}$$

The explicit representation of the LBSDE yields **[Ma, Zhang '02]**

$$Z_t = \nabla Y_t [\nabla X_t]^{-1} \sigma(t, X_t)$$

$$= \mathbb{E}[\Phi'(X_T) \nabla X_T \Gamma_T^t + \int_t^T f'_X(s, X_s, Y_s, Z_s) \nabla X_s \Gamma_T^s ds | \mathcal{F}_t] [\nabla X_t]^{-1} \sigma(t, X_t)$$

Z-regularity

$$\sum_{i=0}^{N-1} \int_{t_i}^{t_{i+1}} \mathbb{E} |Z_t - \bar{Z}_{t_i}|^2 dt$$

Following from this representation, to Ito-decomposition of Z contains :

- ▶ an absolutely continuous part (in dt) \rightarrow **easy to handle.**
- ▶ a martingale part M (in dW_t) :

$$\sum_{i=0}^{N-1} \int_{t_i}^{t_{i+1}} \mathbb{E} |M_t - \bar{M}_{t_i}|^2 dt \leq |\pi| \mathbb{E}(M_T^2 - M_0^2)!!$$

Possible extensions to \mathbb{L}_∞ -functionals [Zhang '04], to jumps [Bouchard, Elie '08], to RBSDE [Bouchard, Chassagneux '06], to BSDE with random terminal time [Bouchard, Moreau '09].

Other methods :Gobet and alii

- ▶ The case of irregular function $\Phi(X_T)$, with strict ellipticity
- ▶ Error expansion for smooth data and uniform grid [G.,Labart '07]
- ▶ Resolution by Picard's iteration, as limit of linear BSDE :
[Bender, Denk '07] ; [G.,Labart '09] with adaptive control variates.
Smaller errors propagation compared to the dynamic programming equation.

Computations of the conditional expectations

Our objective : to implement the dynamic programming equation = to compute the conditional expectations \rightarrow the crucial step !!

Different points of view :

- ▶ the conditional expectation is a projection operator : if $Y \in \mathbb{L}_2$, then

$$\mathbb{E}(Y|X) = \text{Arg} \min_{m \in \mathbb{L}_2(\mathbb{P}^X)} \mathbb{E}(Y - m(X))^2.$$

\rightarrow this is a least-squares problem. What for ?

- To simulate the random variable $m(X)$? one only needs its law.
- To compute the regression function m ? finding a function of dimension $= \dim(X) \rightarrow$ curse of dimensionality.

- ▶ How many regression function to compute?

Answer. For the DPE of BSDEs, N regression functions and $N \rightarrow \infty$.

$$\begin{cases} v^N(t_i, x) = \frac{1}{\Delta t_i} \mathbb{E}(u^N(t_{i+1}, X_{t_{i+1}}^N) \Delta W_{t_i}^N = x), \\ u^N(t_i, x) = \mathbb{E}(u^N(t_{i+1}, X_{t_{i+1}}^N) + \Delta t_i f(t_i, x, u^N(t_{i+1}, X_{t_{i+1}}^N), v^N(t_{i+1}, x)) | X_{t_i}^N = x) \\ u^N(T, x) = \Phi(x). \end{cases}$$

- ▶ In which points $X \in \mathbb{R}^d$?

Answer. Potentially, many ...

**All is a question of global efficiency = balance
between accuracy and computational cost**

Markovian setting

Based on $\mathbb{E}(g(X_{t_{i+1}})|X_{t_i}) = \int g(x)\mathbb{P}_{X_{t_{i+1}}|X_{t_i}}(dx) = m(X_{t_i})$.

If $m(\cdot)$ are required at only few values of $X_{t_i} = x_1, \dots, x_n$:

- ▶ one can simulate M independant paths of $X_{t_{i+1}}$ starting from $X_{t_i} = x_1, \dots, x_n$ and average them out (usual Monte Carlo procedures).
- ▶ but if needed for many i , exponentially growing tree !!

How to put constraints on the complexity ?

One possibility for one-dimensional BM (or Geometric BM) : replace the true dynamics by that of a Bernoulli random walk (**binomial tree**).

3.2 Representation of conditional expectation using Malliavin calculus

[Fournié, Lasry, Lebuchoux, Lions '01 ; Bouchard, Touzi '04 ; Bally, Caramellino, Zanette '05 ...]

Theorem. [integration by parts formula] Suppose that for any smooth f , one has

$$\mathbb{E}(f^k(F)G) = \mathbb{E}(f(F)H_k(F, G))$$

for some r.v. $H_k(F, G)$, depending on F, G , on the multi-index k but not on f .

Then, one has

$$\mathbb{E}(G|F = x) = \frac{\mathbb{E}(\mathbf{1}_{F_1 \leq x_1, \dots, F_d \leq x_d} H_{1, \dots, 1}(F, G))}{\mathbb{E}(\mathbf{1}_{F_1 \leq x_1, \dots, F_d \leq x_d} H_{1, \dots, 1}(F, 1))}.$$

Formal proof ($d = 1$) :

- ▶ The H are obtained using Malliavin calculus, or a direct integration by parts when densities are known.
- ▶ Actually, we look for $H(F, G) = G\tilde{H}(F, G)$. Representation with factorization not so immediate to obtain (possible for SDE).
- ▶ In practice, large variance \rightarrow need some extra localization procedures.
- ▶ For non trivial dynamics, the computational time needed to simulate H may be high.
- ▶ For BSDEs, available rates of convergence w.r.t. N and M **[Bouchard, Touzi '04]** using N independent set of simulated paths.

Statistical regression model : $Y = m(X) + \epsilon$, with $\mathbb{E}[\epsilon|X] = 0$.

X is called the (random) design.

Large literature on statistical tools to approximate $\mathbb{E}[Y|X]$.

References [Hardle '92 ; Bosq, Lecoutre '87 ; Gyorfi, Kohler, Krzyzak, Walk '02]

Problem : compute $m(\cdot)$ using M independent (?) samples $(Y_i, X_i)_{1 \leq i \leq M}$.

Usually estimation errors in the literature are not sufficient for our purpose :

- ▶ the law X may not have a density w.r.t. Lebesgue measure.
- ▶ the support of the law of the X is never bounded !!

Discussions of non parametric regression tools from theoretical/practical points of view

3.3.1. Kernel estimators

$$\mathbb{E}[Y|W = x] \approx \frac{\frac{1}{h^d} \sum_{i=1}^M K\left(\frac{x-X_i}{h}\right) Y_i}{\frac{1}{h^d} \sum_{i=1}^M K\left(\frac{x-X_i}{h}\right)} = m_{M,h}(x), \text{ where}$$

- ▶ the kernel function is defined on the compact support $[-1, 1]$, bounded, even, non-negative, C_p^2 and $\int_{\mathbb{R}^d} K(u) du = 1$,
- ▶ $h > 0$ is the bandwidth.

Non-integrated \mathbb{L}_2 -error estimates available.

3.3.1. Projection on a set of functions

Set of functions : $(\phi_k)_{0 \leq k \leq K}$.

$$\mathbb{E}(Y|X) = \arg \min_g \mathbb{E}(Y - g(X))^2 \approx \arg \min_{\sum_{k=1}^K \alpha_k \phi_k(\cdot)} (Y - \sum_{k=1}^K \alpha_k \phi_k(X))^2.$$

Computations of the optimal coefficients $(\alpha_k)_k$: it solves the normal equation

$$A\alpha = \mathbb{E}(Y\phi), \text{ where } A_{i,j} = \mathbb{E}(\phi_i(X)\phi_j(X)), [\mathbb{E}(Y\phi)]_i = \mathbb{E}(Y\phi_i(X)).$$

- For simplicity, one should have a system of orthonormal functions (w.r.t the law of X).

- ▶ If the system is not orthonormal, one should compute A and invert it.

Its dimensions is expected to be very large : $K \rightarrow \infty$ to ensure convergent approximations.

Presumably big instabilities (ill-conditioned matrix) to solve this least-squares problem **[Golub, Van Loan '96]**.

- ▶ In practice, A is computed using simulations, as well $\mathbb{E}[Y\phi]$.

Equivalent to solve the **empirical least-squares problem** :

$$(\alpha_k^M)_k = \arg \min_{\alpha} \frac{1}{M} \sum_{m=1}^M (Y^m - \sum_{k=1}^K \alpha_k \phi_k(X^m))^2.$$

CLT At fixed K , if A is invertible, one has

$$\lim_{M \rightarrow \infty} \sqrt{M}(\alpha^M - \alpha) = \mathcal{N}(0, \dots).$$

The case of polynomial functions

- ▶ Popular choice.
- ▶ Smooth approximation.
- ▶ Global approximation : within few polynomials, a smooth $m(\cdot)$ can be very well approximated.
- ▶ But show convergence for non smooth functions (non-linear BSDEs may lead non-smooth functions).
- ▶ Do projections on polynomials converge to $m(\cdot)$?

$\bigoplus_{k \geq 0} (P)_k(X) = \mathbb{L}_2(X)$? If for some $a > 0$ one has $\mathbb{E}(e^{a|X|}) < \infty$, then polynomials are dense in \mathbb{L}_2 -functions.

Proof. Related to the moment problems. Is a r.v. characterized by its polynomial moment ? In particular, if X is log-normal, orthonomials of X are not dense in \mathbb{L}_2 (Feller counter-exemple) !!

The case of local approximation

Piecewise constant approximations $\phi_k = \mathbf{1}_{\mathcal{C}_k}$, where the subsets $(\mathcal{C}_k)_k$ forms a tessellation of a part of \mathbb{R}^d : $\mathcal{C}_k \cap \mathcal{C}_l = \emptyset$ for $l \neq k$.

$$\arg \inf_{g = \sum_k \alpha_k \mathbf{1}_{\mathcal{C}_k}} \mathbb{E}(Y - g(X))^2 \text{ or } \arg \inf_{g = \sum_k \alpha_k \mathbf{1}_{\mathcal{C}_k}} \mathbb{E}^M(Y - g(X))^2?$$

The "matrix"

$A = (\mathbb{E}(\phi_i(X)\phi_j(X)))_{i,j}$ is diagonal : $A = \text{Diag}(\mathbb{P}(X \in \mathcal{C}_i)_i) \implies$

$$\alpha_k = \begin{cases} \frac{\mathbb{E}(Y \mathbf{1}_{X \in \mathcal{C}_k})}{\mathbb{P}(X \in \mathcal{C}_k)} = \mathbb{E}(Y | X \in \mathcal{C}_k) & \text{if } \mathbb{P}(X \in \mathcal{C}_k) > 0, \\ 0 & \text{if } \mathbb{P}(X \in \mathcal{C}_k) = 0, \end{cases}$$

$$M = \left[\frac{1}{\#\{m: X^m \in \mathcal{C}_k\}} \sum_{m: X^m \in \mathcal{C}_k} Y^m \right]_{k=1, \dots, K} \text{ if } \#\{m: X^m \in \mathcal{C}_k\} \geq 0,$$

Rate of approximations of a Lipschitz regression function $m(\cdot)$

Size of the tessellation : $|\mathcal{C}| \leq \sup_I \sup_{(x,y) \in \mathcal{C}_I} |x - y|$.

Given a probability measure $\mu : \mu = \mathbb{P}_X$ or $\mu = \frac{1}{M} \sum_{m=1}^M \delta_{X^m}(\cdot)$.

$$\begin{aligned}
 & \inf_{g = \sum_k \alpha_k \mathbf{1}_{\mathcal{C}_k}} \int_{\mathbb{R}^d} |g(x) - m(x)|^2 \mu(dx) \\
 & \leq \sum_k \int_{\mathcal{C}_k} |m(x_k) - m(x)|^2 \mu(dx) + \int_{[\cup_k \mathcal{C}_k]^c} m^2(x) \mu(dx) \\
 & \leq \sum_k |\mathcal{C}|^2 \mu(\mathcal{C}_k) + |m|_\infty^2 \mu([\cup_k \mathcal{C}_k]^c) \leq |\mathcal{C}|^2 + |m|_\infty^2 \mu([\cup_k \mathcal{C}_k]^c).
 \end{aligned}$$

► We expect the tessellation size to be small.

► The complementary $\mu([\cup_k \mathcal{C}_k]^c)$ has to be small (tail estimate).

Efficient choice of tessellation ?

Given $x \in \mathbb{R}^d$, how to locate efficiently the \mathcal{C}_k such that $x \in \mathcal{C}_k$?

- **Voronoi tessellations** associated to a sample $(X^K)_{1 \leq k \leq K}$ of the underlying r.v. $X : \mathcal{C}_k = \{z \in \mathbb{R}^d : |z - X^k| = \min_l |z - X^l|\}$.
Closed to quantization ideas.

Theoretically, there exists searching algorithms with a cost $\mathcal{O}(\log(K))$.

- **Regular grid (hepercubes).**

$k = (k_1, \dots, k_d) \in \{0, \dots, K_1 - 1\} \times \dots \times \{0, \dots, K_d - 1\}$ define

$$\mathcal{C}_k = [-x_{1,\min} + \Delta x_1 k_1, -x_{1,\min} + \Delta x_1 (k_1 + 1)] \times \dots \times [-x_{d,\min} + \Delta x_d k_d, -x_{d,\min} + \Delta x_d (k_d + 1)]$$

Tessellation size = $\mathcal{O}(\max_i \Delta x_i)$

Quick search formula :

3.4 Model-free estimation of the regression error [GKKW02]

In the BSDEs framework, see [Lemor, G., Warin '06].

Working assumptions :

- ▶ $Y = m(X) + \epsilon$ with $\mathbb{E}(\epsilon|X) = 0$.
- ▶ Data : sample of independant copies $(X_1, Y_1), \dots, (X_n, Y_n)$.
- ▶ $\sigma^2 = \sup_x \text{Var}(Y|X = x) < \infty$
- ▶ $F_n = \text{Span}(f_1, \dots, f_{K_n})$ a linear vector space of dimension K_n , which may depend on the data !

Notations : $|f|_n^2 = \frac{1}{n} \sum_{i=1}^n f^2(X_i)$. Write μ^n for the empirical measure associated to (X_1, \dots, X_n) .

$$\hat{m}_n(.) = \arg \min_{f \in F_n} \frac{1}{n} \sum_{i=1}^n |f(X_i) - Y_i|^2.$$

Proof

W.l.o.g., we can assume that

- ▶ (f_1, \dots, f_{K_n}) is orthonormal family in $\mathbb{L}_2(\mu^n)$: $\frac{1}{n} \sum_i f_k(X_i) f_l(X_i) = \delta_{k,l}$.

\implies The solution of $\arg \min_{f \in F_n} \frac{1}{n} \sum_{i=1}^n |f(X_i) - Y_i|^2$ is given by

$$\hat{m}_n(\cdot) = \sum_j \alpha_j f_j(\cdot) \text{ with } \alpha_j = \frac{1}{n} \sum_i f_j(X_i) Y_i.$$

Lemma. Denote $\mathbb{E}^*(\cdot) = \mathbb{E}(\cdot | X_1, \dots, X_n)$. Then $\mathbb{E}^*(\tilde{m}_n(\cdot))$ is the least-squares solution of $\arg \min_{f \in F_n} \frac{1}{n} \sum_{i=1}^n |f(X_i) - m(X_i)|^2 = \arg \min_{f \in F_n} |f - m|_n^2$.

Proof.

- ▶ The above least-squares solution is given by $\sum_j \alpha_j^* f_j(\cdot)$ with

$$\alpha_j^* = \frac{1}{n} \sum_i f_j(X_i) m(X_i).$$

Pythagore theorem : $|\tilde{m}_n - m|^2 = |\tilde{m}_n - \mathbb{E}^*(\tilde{m}_n)|_n^2 + |\mathbb{E}^*(\tilde{m}_n) - m|_n^2$.

$$\begin{aligned} \text{Then, } \mathbb{E}^*|\tilde{m}_n - m|_n^2 &= \mathbb{E}^*|\tilde{m}_n - \mathbb{E}^*(\tilde{m}_n)|_n^2 + |\mathbb{E}^*(\tilde{m}_n) - m|_n^2 \\ &= \mathbb{E}^*|\tilde{m}_n - \mathbb{E}^*(\tilde{m}_n)|_n^2 + \min_{f \in F_n} |f - m|_n^2. \end{aligned}$$

Since $(f_j)_j$ is orthonormal in $\mathbb{L}_2(\mu_n)$, we have

$$|\tilde{m}_n - \mathbb{E}^*(\tilde{m}_n)|_n^2 = \sum_j |\alpha_j - \mathbb{E}^*(\alpha_j)|^2.$$

Thus, using $\alpha_j - \mathbb{E}^*(\alpha_j) = \frac{1}{n} \sum_i f_j(X_i)(Y_i - m(X_i))$, we have

$$\begin{aligned} \mathbb{E}^*|\tilde{m}_n - \mathbb{E}^*(\tilde{m}_n)|_n^2 &= \sum_j \frac{1}{n^2} \mathbb{E}^* \sum_{i,l} f_j(X_i) f_j(X_l) (Y_i - m(X_i))(Y_l - m(X_l)) \\ &= \sum_j \frac{1}{n^2} \sum_i f_j^2(X_i) \text{Var}(Y_i|X_i) \end{aligned}$$

since **the $(\epsilon_i)_i$ conditionnaly on (X_1, \dots, X_n) are centered.**

$$\implies \mathbb{E}^*|\tilde{m}_n - \mathbb{E}^*(\tilde{m}_n)|_n^2 \leq \sigma^2 \sum \frac{1}{n^2} \sum f_j^2(X_i) = \sigma^2 \frac{K_n}{n}.$$

Uniform law of large numbers

$Z_{1:n} = (Z_1, \dots, Z_n)$ a i.i.d. sample of size n .

For $\mathcal{G} \subset \{g : \mathbb{R}^d \mapsto [0, B]\}$, one needs to quantify

$$\mathbb{P}[\forall g \in \mathcal{G} : |\frac{1}{n} \sum_{i=1}^n g(Z_i) - \mathbb{E}g(Z)| > \epsilon]$$

as a function of ϵ and n ?

By Borel-Cantelli lemma, may lead to uniform laws of large numbers :

$$\sup_{g \in \mathcal{G}} |\frac{1}{n} \sum_{i=1}^n g(Z_i) - \mathbb{E}g(Z)| \rightarrow 0 \text{ a.s.}$$

ϵ -cover of \mathcal{G}

Definition. For a class of functions \mathcal{G} and a given empirical measure μ^n associated to n points $Z_{1:n} = (Z_1, \dots, Z_n)$, we define a ϵ -cover of \mathcal{G} w.r.t. $\mathbb{L}_1(\mu^n)$ by a collection (g_1, \dots, g_N) in \mathcal{G} such that

for any $g \in \mathcal{G}$, there is a $j \in \{1, \dots, N\}$ s.t. $|g - g_j|_{\mathbb{L}_1(\mu^n)} < \epsilon$.

Set $\mathcal{N}_1(\epsilon, \mathcal{G}, \mathbf{Z}_{1:n})$ = the smallest size N of ϵ -cover of \mathcal{G} w.r.t. $\mathbb{L}_1(\mu^n)$.

Theorem. For $\mathcal{G} \subset \{g : \mathbb{R}^d \mapsto [-B, B]\}$. For any n and any $\epsilon > 0$, one has

$$\mathbb{P}(\forall g \in \mathcal{G} : |\frac{1}{n} \sum_{i=1}^n g(\mathbf{Z}_i) - \mathbb{E}g(\mathbf{Z})| > \epsilon) \leq 8\mathbb{E}(\mathcal{N}_1(\epsilon/8, \mathcal{G}, \mathbf{Z}_{1:n})) \exp(-\frac{n\epsilon^2}{512B^2}).$$

Theorem. If $\mathcal{G} = \{-B \vee \sum_k \alpha_k \phi_k(\cdot) \wedge B : (\alpha_1, \dots, \alpha_K) \in \mathbb{R}^K\}$, then

$$\mathcal{N}_1(\epsilon, \mathcal{G}, \mathbf{Z}_{1:n}) \leq 3 \left(\frac{4eB}{\epsilon} \log \left(\frac{4eB}{\epsilon} \right) \right)^{K+1}$$

3.5 Applications to numerical solution of BSDEs using empirical simulations [LGW06]

Regular time grid with time step $h = \frac{T}{N}$ + Lipschitz f , Φ , b and σ .

Towards an approximation of the regression operators

Truncation of the tails using a threshold $R = (R_0, \dots, R_d)$:

$$[\Delta W_{l,k}]_w = (-R_0 \sqrt{h}) \vee \Delta W_{l,k} \wedge (R_0 \sqrt{h}),$$

$$f^R(t, x, y, z) = f(t, -R_1 \vee x_1 \wedge R_1, \dots, -R_d \vee x_d \wedge R_d, y, z),$$

$$\Phi^R(x) = \Phi(-R_1 \vee x_1 \wedge R_1, \dots, -R_d \vee x_d \wedge R_d).$$

→ Localized BSDEs

Define $Y_T^{N,R}(X_{t_k}^N) = \Phi^R(X_{t_k}^N)$ and

Proposition. For some **Lipschitz** functions $y_k^{N,R}(\bullet)$ and $z_k^{N,R}(\bullet)$, one has :

$$\begin{cases} Z_{l,t_k}^{N,R} &= \frac{1}{h} \mathbb{E}(Y_{t_{k+1}}^{N,R} [\Delta W_{l,k}]_\omega | \mathcal{F}_{t_k}) = z_{l,k}^{N,R}(X_{t_k}^N). \\ Y_{t_k}^{N,R} &= \mathbb{E}(Y_{t_{k+1}}^{N,R} + hf^R(t_k, X_{t_k}^N, Y_{t_{k+1}}^{N,R}, Z_{t_k}^{N,R}) | \mathcal{F}_{t_k}) = y_k^{N,R}(X_{t_k}^N). \end{cases}$$

a) The Lipschitz constants of $y_k^{N,R}(\bullet)$ and $N^{-1/2} z_k^{N,R}(\bullet)$ are uniform in N and R .

b) **Bounded functions** : $\sup_N (\|y_k^{N,R}(\bullet)\|_\infty + N^{-1/2} \|z_k^{N,R}(\bullet)\|_\infty) = C_\star < \infty$

Proposition. (Convergence as $|R| \uparrow \infty$) For h small enough, one has

$$\begin{aligned} & \max_{0 \leq k \leq N} \mathbb{E} |Y_{t_k}^{N,R} - Y_{t_k}^N|^2 + h \mathbb{E} \sum_{k=0}^{N-1} |Z_{t_k}^{N,R} - Z_{t_k}^N|^2 \\ & \leq C \mathbb{E} |\Phi(X_{t_n}^N) - \Phi^R(X_{t_n}^N)|^2 + C \frac{1+R^2}{h} \sum_{k=0}^{N-1} \mathbb{E} (|\Delta W_k|^2 \mathbf{1}_{|\Delta W_k| \geq R_0 \sqrt{h}}) \\ & + Ch \mathbb{E} \sum_{k=0}^{N-1} |f(t_k, X_{t_k}^N, Y_{t_{k+1}}^N, Z_{t_k}^N) - f^R(t_k, X_{t_k}^N, Y_{t_{k+1}}^N, Z_{t_k}^N)|^2. \end{aligned}$$

Approximation of $y_k^{N,R}(\bullet)$ and $z_k^{N,R}(\bullet)$

Projection on a finite dimensional space :

$$y_k^{N,R}(\bullet) \approx \alpha_{0,k} \cdot p_{0,k}(\bullet), \quad z_{l,k}^{N,R}(\bullet) \approx \alpha_{l,k} \cdot p_{l,k}(\bullet).$$

(for instance, hypercubes as presented before).

Coefficients will be computed by extra M independent simulations of $(X_{t_k}^N)_k$ and $(\Delta W_k)_k \rightarrow \{(X_{t_k}^{N,m})_k\}_m$ and $\{(\Delta W_k^m)_k\}_m$ (**only one set of simulated paths**).

In addition, we impose **boundedness properties** :

$$y_k^{N,R,M}(\bullet) = [\alpha_{0,k}^M \cdot p_{0,k}(\bullet)]_y, \quad z_{l,k}^{N,R,M}(\bullet) \approx [\alpha_{l,k}^M \cdot p_{l,k}(\bullet)]_z,$$

The final algorithm

→ Initialization : for $k = N$ take $y_N^{N,R}(\cdot) = \Phi^R(\cdot)$.

→ Iteration : for $k = N - 1, \dots, 0$, solve the q least-squares problems :

$$\alpha_{l,k}^M = \arg \inf_{\alpha} \frac{1}{M} \sum_{m=1}^M |y_{k+1}^{N,R,M}(X_{t_{k+1}}^{N,m}) \frac{[\Delta W_{l,k}^m]_{\omega}}{h} - \alpha \cdot p_{l,k}(X_{t_k}^{N,m})|^2$$

Then compute $\alpha_{0,k}^M$ as the minimizer of

$$\sum_{m=1}^M |y_{k+1}^{N,R,M}(X_{t_{k+1}}^{N,m}) + h f^R(t_k, X_{t_k}^{N,m}, y_{k+1}^{N,R,M}(X_{t_{k+1}}^{N,m}), [\alpha_{l,k}^M \cdot p_{l,k}(X_{t_k}^{N,m})]_z) - \alpha \cdot p_{0,k}(X_{t_k}^{N,m})|^2$$

Then define $y_k^{N,R,M}(\bullet) = [\alpha_{0,k}^M \cdot p_{0,k}(\bullet)]_y$, $z_{l,k}^{N,R,M}(\bullet) = [\alpha_{l,k}^M \cdot p_{l,k}(\bullet)]_z$.

Error analysis

Robust error bounds

Theorem. Under Lipschitz conditions (only!), one has

$$\begin{aligned}
 & \max_{0 \leq k \leq N} \mathbb{E} |Y_{t_k}^{N,R} - y_k^{N,R,M}(S_{t_k}^N)|^2 + h \sum_{k=0}^{N-1} \mathbb{E} |Z_{t_k}^{N,R} - z_k^{N,R,M}(S_{t_k}^N)|^2 \\
 & \leq C \frac{C_\star^2 \log(M)}{M} \sum_{k=0}^{N-1} \sum_{l=0}^q \mathbb{E}(K_{l,k}^M) + Ch \\
 & + C \sum_{k=0}^{N-1} \left\{ \inf_{\alpha} \mathbb{E} |y_k^{N,R}(S_{t_k}^N) - \alpha \cdot p_{0,k}(S_{t_k}^N)|^2 + \sum_{l=1}^q \left\{ \inf_{\alpha} \mathbb{E} |\sqrt{h} z_{l,k}^{N,R}(S_{t_k}^N) - \alpha \cdot p_{l,k}(S_{t_k}^N)|^2 \right\} \right\} \\
 & + C \frac{C_\star^2}{h} \sum_{k=0}^{N-1} \left\{ \mathbb{E} \left[K_{0,k}^M \exp\left(-\frac{Mh^3}{72C_\star^2 K_{0,k}^M}\right) \exp\left(CK_{0,k+1} \log \frac{CC_\star(K_{0,k}^M)^{\frac{1}{2}}}{h^{\frac{3}{2}}}\right) \right] \right. \\
 & + h \mathbb{E} \left[K_{l,k}^M \exp\left(-\frac{Mh^2}{72C_\star^2 R_0^2 K_{l,k}^M}\right) \exp\left(CK_{0,k+1} \log \frac{CC_\star R_0(K_{l,k}^M)^{\frac{1}{2}}}{h}\right) \right] \\
 & \left. + \exp\left(CK_{0,k} \log \frac{CC_\star}{h^{\frac{3}{2}}}\right) \exp\left(-\frac{Mh^3}{72C_\star^2}\right) \right\}.
 \end{aligned}$$

Convergence of the parameters in the cases of HC functions

For a global squared error of order $\epsilon = \frac{1}{N}$, choose :

- ① Edge of the hypercube : $\delta \sim \frac{C}{N}$.
- ② Number of simulations : $M \sim N^{3+2d}$.

Available for a large class of models on X , which depend essentially on \mathbb{L}_2 bounds on the solution (no ellipticity condition, with or without jump. . .).

Complexity/accuracy

Global complexity : $\mathcal{C} \sim \epsilon^{-\frac{1}{4+2d}}$.

Techniques of **local duplicating of paths** : $\mathcal{C} \sim \epsilon^{-\frac{1}{4+d}}$.

3.6 Numerical results (mainly due to J.P. Lemor)

Ex.1 : bid-ask spread for interest rates

- ▶ Black-Scholes model and $\Phi(\mathbf{S}) = (S_T - K_1)_+ - 2(S_T - K_2)_+$.
- ▶ $f(t, x, y, z) = -\{yr + z\theta - (y - \frac{z}{\sigma})_-(R - r)\}$, $\theta = \frac{\mu - r}{\sigma}$.

- ▶ Parameters :

μ	σ	r	R	T	S_0	K_1	K_2
0.05	0.2	0.01	0.06	0.25	100	95	105

M	$N = 5, \delta = 5$ $D = [60, 140]$	$N = 20, \delta = 1$ $D = [60, 200]$	$N = 50, \delta = 0.5$ $D = [60, 200]$
128	3.05(0.27)	3.71(0.95)	3.69(4.15)
512	2.93(0.11)	3.14(0.16)	3.48(0.54)
2048	2.92(0.05)	3.00(0.03)	3.08(0.12)
8192	2.91(0.03)	2.96(0.02)	2.99(0.02)
32768	2.90(0.01)	2.95 (0.01)	2.98(0.01)

Global polynomials (GP)

Polynomials of d variables with a maximal degree.

M	$N = 5$ $d_y = 1, d_z = 0$	$N = 20$ $d_y = 2, d_z = 1$	$N = 50$ $d_y = 4, d_z = 2$	$N = 50$ $d_y = 9, d_z =$
128	2.87(0.39)	3.01(0.24)	3.02(0.22)	3.49(1.57)
512	2.82(0.20)	2.94(0.12)	2.97(0.09)	3.02(0.1)
2048	2.78(0.07)	2.92(0.07)	2.92(0.04)	2.97(0.03)
8192	2.78(0.05)	2.92(0.04)	2.92(0.02)	2.96(0.01)
32768	2.79(0.03)	2.91(0.02)	2.91(0.01)	2.95(0.01)

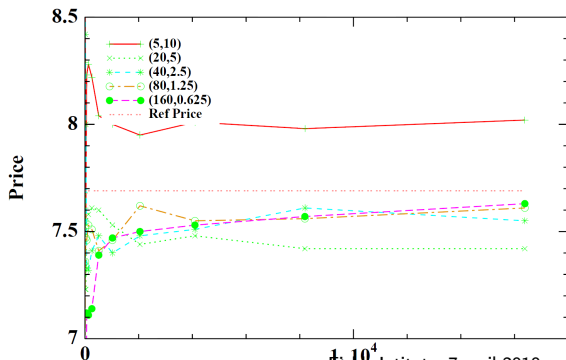
Table: Results for the calls combination using **GP**.

Ex.2 : locally-risk minimizing strategies (FS decomposition)

Heston stochastic volatility models [Heath,Platen,Schweizer '02] :

$$\frac{dS_t}{S_t} = \gamma Y_t^2 dt + Y_t dW_t, \quad dY_t = \left(\frac{c_0}{Y_t} - c_1 Y_t \right) dt + c_2 dB_t.$$

Functions **HC**,
parameters (N, δ) .



American options via RBDSEs : several approaches

1. Talking the **max** with obstacle \rightarrow Bermuda options (**lower approximation**)

$$Y_{t_k}^n = \max(\Phi(t_k, S_{t_k}^N), \mathbb{E}(Y_{t_{k+1}}^N | \mathcal{F}_{t_k}) + hf(t_k, S_{t_k}^N, Y_{t_k}^N, Z_{t_k}^N)),$$

$$Z_{l,t_k}^N = \frac{1}{h} \mathbb{E}(Y_{t_{k+1}}^N \Delta W_{l,k} | \mathcal{F}_{t_k}).$$

2. **Penalization**. Obtained as the limit of standard BSDEs with driver $f(s, S_s, Y_s, Z_s) + \lambda(Y_s - \Phi(s, S_s))_-$ with $\lambda \uparrow +\infty$.

Lower approximation.

3. **Regularization** of the increasing process : when

$$d\Phi(t, S_t) = U_t dt + V_t dW_t + dA_t^+,$$

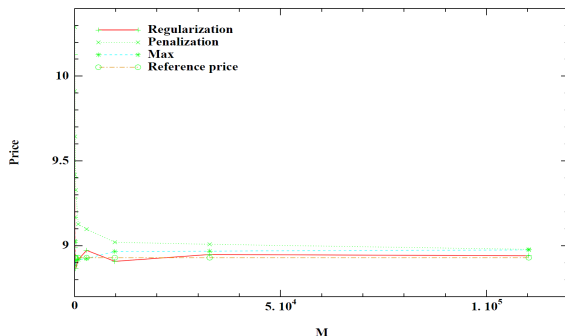
Ex.3 : American options on tree assets

- ▶ Payoff $g(x) = (K - (\prod_{i=1}^3 x_i)^{\frac{1}{3}})^+$.

- ▶ Black-Scholes parameters :

T	r	σ	K	S_0^i	d
1	0.05	0.41d	100	100	1

- ▶ Reference price **8.93** (PDE method).



Functions **HC(1,0)** with local polynomials of degree 1 for Y and 0 for Z .

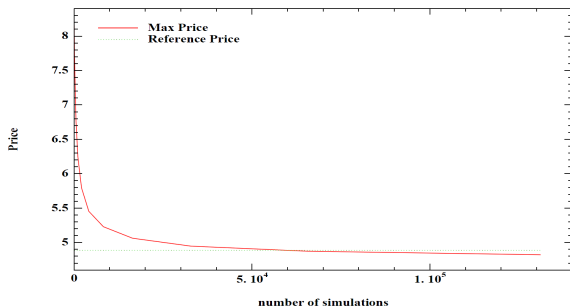
Regularisation : $N = 32$,
 $\delta = 9$, $\lambda = 2$.

Max : $N = 44$, $\delta = 7$.

Penalization : $N = 60$,
 $\delta = 2$, $\lambda = 2$.

Ex.4 : American options on ten assets

- ▶ $d = 10 = 2p$. Multidimensional Black-Scholes model :
$$\frac{dS_t^l}{S_t^l} = (r - \mu_l)dt + \sigma_l dW_t^l.$$
- ▶ Payoff : $\max(x_1 \cdots x_p - x_{p+1} \cdots x_{2p}, 0)$.
- ▶ $r = 0$, dividend rate $\mu_1 = -0.05$, $\mu_l = 0$ for $l \geq 2$. $\sigma_l = \frac{0.2}{\sqrt{d}}$.
 $T = 0.5$. $S_0^i = 40^{\frac{2}{d}}$, $1 \leq i \leq p$. $S_0^i = 36^{\frac{2}{d}}$, $p+1 \leq i \leq 2p$.
- ▶ Reference price **4.896**, obtained with a PDE method
[Villeneuve, Zanette 2002].
- ▶ Price with quantization algorithm : 4.9945
[Bally-Pages-Printemps 2005].



Functions **HC(1,0)**.

Max : $N = 60$, $\delta = 0.6$.

Computational time :
15 seconds.

References