

Diffusion processes and coalescent trees

Bob Griffiths

University of Oxford

The frequency of a mutation

The relative frequency $\{X(t), t \geq 0\}$ of genes of type a in a population of two types a and A is modelled by a diffusion process with generator

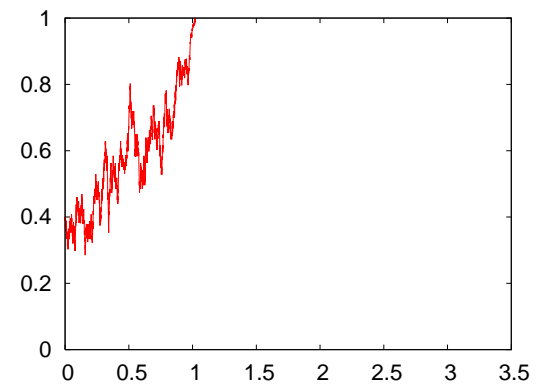
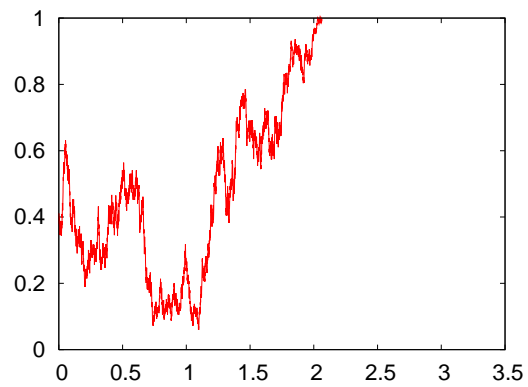
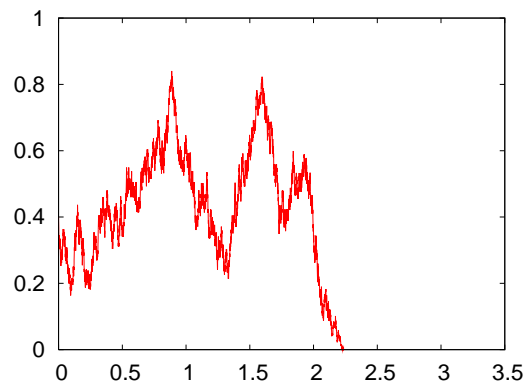
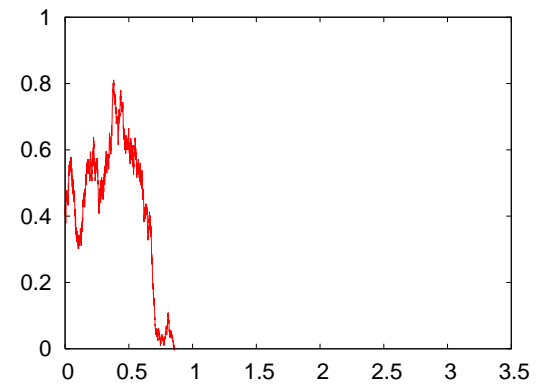
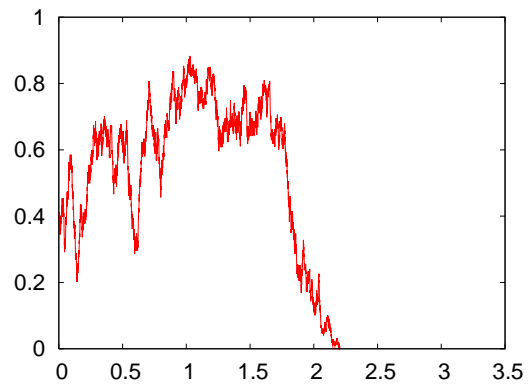
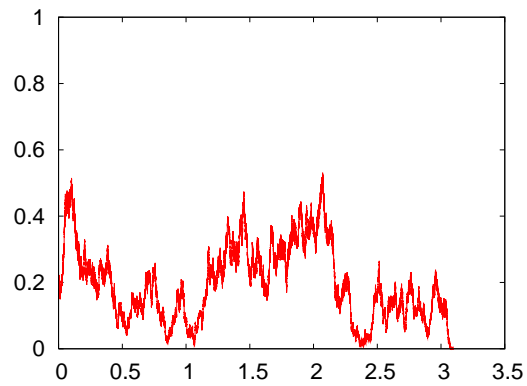
$$L = \frac{1}{2}x(1-x)\frac{\partial^2}{\partial x^2}$$

Simplest model with no mutation or selection.

Denote $\Delta X(t) = X(t + \Delta t) - X(t)$

$$\begin{aligned} E(\Delta X(t) \mid X(t) = x) &= 0\Delta t + o(\Delta t) \\ \text{Var}(\Delta X(t) \mid X(t) = x) &= x(1-x)\Delta t + o(\Delta t) \end{aligned}$$

Kimura (1955), Bochner (1954) in a different context.



$X(t)$, Neutral model, no mutation

Wright-Fisher model N genes, Types a, A . Fixed population size of N genes. Offspring in a generation choose their parents at random with replacement from the prior generation and inherit their type.

$X^N(\tau), \tau = 0, 1, \dots$ is the relative frequency of a genes in generation τ .

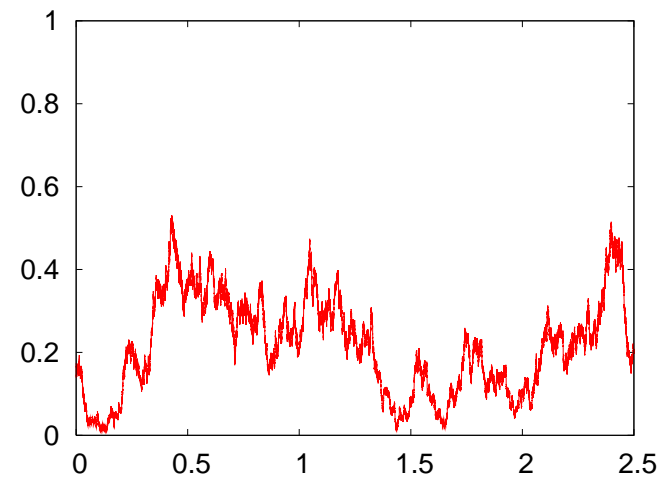
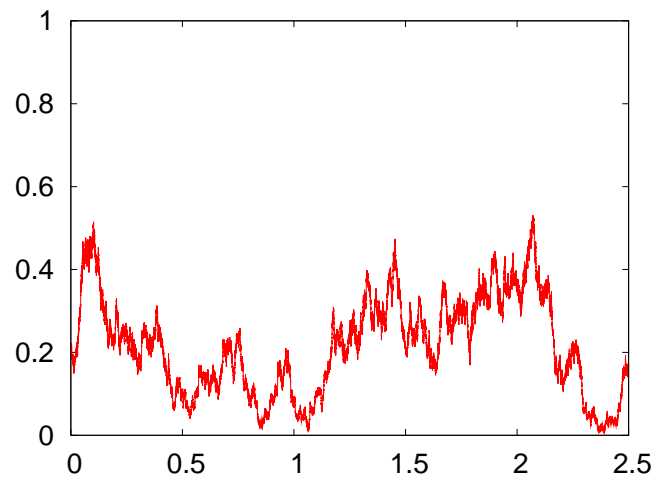
$$E(\Delta X^N(\tau) \mid X^N(\tau) = x) = 0$$

$$\text{Var}(\Delta X^N(\tau) \mid X^N(\tau) = x) = x(1 - x) \cdot \frac{1}{N}$$

Higher moments of $\Delta X^N(\tau)$ are of order smaller than $\frac{1}{N}$.

Measure time in units of N generations, so $\delta t = \frac{1}{N}$, then

$$\{X^N([Nt]), t \geq 0\} \implies \{X(t), t \geq 0\}$$



Reversibility of $X(t)$

Transition density in the neutral model

$$f(p, x; t) = x^{-1}(1-x)^{-1} \sum_{i=2}^{\infty} e^{-\frac{1}{2}i(i-1)t} P_i(p) P_i(x)$$

where $\{P_i(\cdot)\}$ are scaled Gegenbauer polynomials, orthogonal on

$$m(x) = x^{-1}(1-x)^{-1}$$

Reversibility before hitting 0 or 1

$$m(x)f(x, p; t) = m(p)f(p, x; t)$$

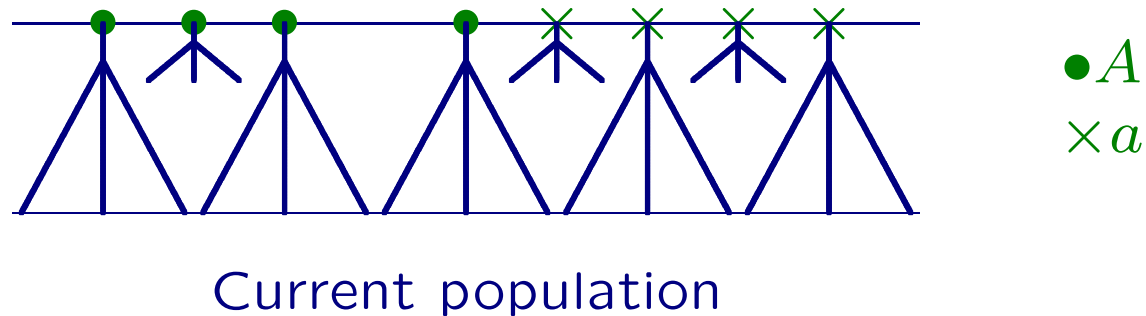
Genealogical form of the transition density

$$f(p, x; t) = \sum_{\ell=1}^{\infty} q_{\ell}(t) \sum_{k=1}^{\ell} \binom{\ell}{k} p^k (1-p)^{\ell-k} \times \frac{\Gamma(\ell)}{\Gamma(k)\Gamma(\ell-k)} x^{k-1} (1-x)^{\ell-k-1}$$

where $\{q_{\ell}(t)\}$ are transition probabilities of a death process (coalescent process) $\{A(t), t \geq 0\}$ with death rates $\mu_k = \binom{k}{2}$ and $A(0) = \infty$.

G (1980), Ethier and G (1993).

Time t back, $A(t) = \ell$ ancestors



The relative frequency of family sizes in the ℓ lines is $\text{Dirichlet}(1, 1, \dots, 1)$.

Line types are chosen from a, A independently with probability $p, 1 - p$.

If k of the ℓ lines are of type a , then the density of a is

$$\frac{\Gamma(\ell)}{\Gamma(k)\Gamma(\ell - k)} x^{k-1} (1 - x)^{\ell-k-1}$$

Two-allele models with mutation

Wright-Fisher model with genes of type a , A

Mutation $a \rightarrow A$ with probability u per gene per generation

Mutation $A \rightarrow a$ with probability v per gene per generation

Diffusion process $\{X(t), t \geq 0\}$ for the relative frequency of a genes at time t .

Generator

$$L = \frac{1}{2}x(1-x)\frac{\partial^2}{\partial x^2} + \frac{1}{2}(-\alpha x + \beta(1-x))\frac{\partial}{\partial x}$$

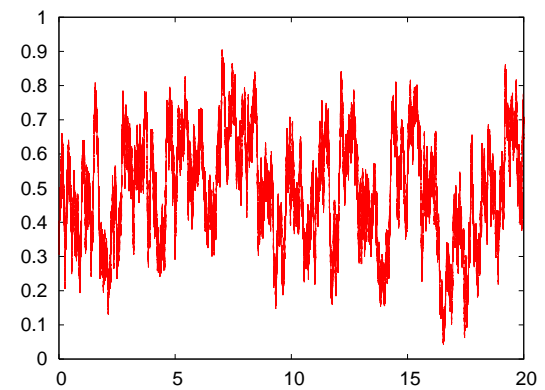
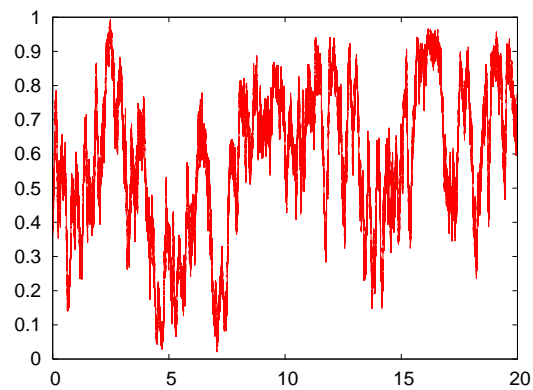
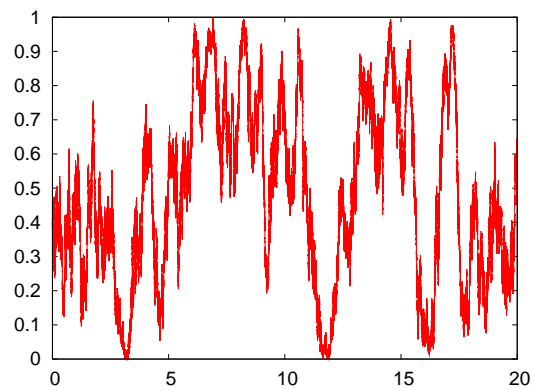
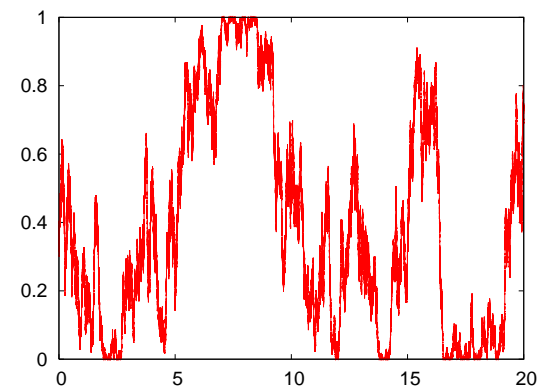
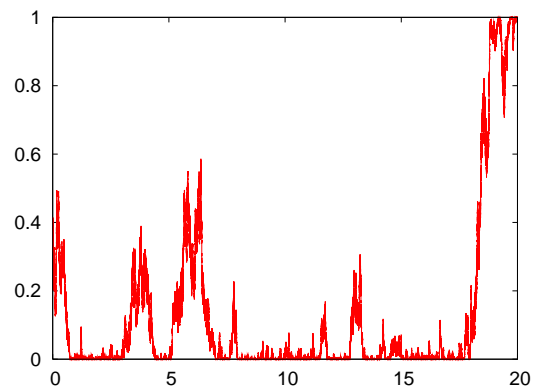
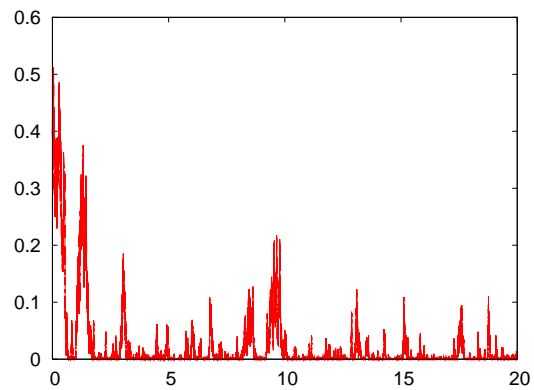
where $\alpha = 2Nu$, $\beta = 2Nv$.

The stationary distribution of the diffusion process is Beta

$$\phi(x; \alpha, \beta) = B(\alpha, \beta)^{-1} x^{\alpha-1} (1-x)^{\beta-1}, \quad 0 < x < 1$$

satisfying the forward equation

$$\frac{1}{2} \frac{\partial^2}{\partial x^2} \left\{ x(1-x) \phi(x; \alpha, \beta) \right\} - \frac{\partial}{\partial x} \left\{ \mu(x) \phi(x; \alpha, \beta) \right\} = 0$$



$X(t)$, Neutral model, with mutation, $\alpha = \beta = 0.01, 0.1, 0.5, 1.0, 2.0, 5.0$

Wright-Fisher diffusion process with d types

$X(t) = (X_1(t), \dots, X_d(t))$ are gene frequencies of d types, labelled $1, 2, \dots, d$ at time $t \geq 0$.

$|X(t)| = X_1(t) + \dots + X_d(t) = 1$.

Mutations $i \rightarrow j$ occur at rate $\frac{1}{2}\epsilon_j$, $i, j = 1, \dots, d$.

Backwards generator

$$\mathcal{L} = \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d x_i (\delta_{ij} - x_j) \frac{\partial^2}{\partial x_i \partial x_j} + \frac{1}{2} \sum_{i=1}^d (\epsilon_i - |\epsilon| x_i) \frac{\partial}{\partial x_i}$$

Stationary distribution is Dirichlet

$$\mathcal{D}(x, \epsilon) = \frac{\Gamma(|\epsilon|)}{\Gamma(\epsilon_1) \cdots \Gamma(\epsilon_d)} x_1^{\epsilon_1-1} \cdots x_d^{\epsilon_d-1}$$

for $x_1, \dots, x_d > 0$ and $|x| = \sum_1^d x_i = 1$

Transition density ($x \rightarrow y$ in time t) is

$$f(x, y; t) = \mathcal{D}(y, \epsilon) \left\{ 1 + \sum_{\{n; n \in \mathbb{Z}_+^{d-1}, |n| > 0\}} \rho_{|n|}(t) Q_n^\circ(x) Q_n^\circ(y) \right\}$$

Eigenfunctions $Q_n^\circ(x)$ are orthonormal polynomials on the Dirichlet distribution.

Eigenvalues are $\rho_{|n|}(t) = \exp\{-\frac{1}{2}|n|(|n| + |\epsilon| - 1)t\}$ repeated $\binom{|n|+d-2}{|n|}$ times corresponding to eigenvectors $Q_n^\circ(x)$, G (1979).

Karlin and McGregor (1975) - Moran model version whose stationary distribution is the Dirichlet-Multinomial. Polynomials are Hahn multivariable analogues explicitly constructed.

Backward generator

$$\mathcal{L} = \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d x_i (\delta_{ij} - x_j) \frac{\partial^2}{\partial x_i \partial x_j} + \frac{1}{2} \sum_{i=1}^d (\epsilon_i - |\epsilon| x_i) \frac{\partial}{\partial x_i}$$

Forward generator

$$\tilde{\mathcal{L}} = \frac{1}{2} \sum_{i=1}^d \sum_{j=1}^d \frac{\partial^2}{\partial x_i \partial x_j} x_i (\delta_{ij} - x_j) - \frac{1}{2} \sum_{i=1}^d \frac{\partial}{\partial x_i} (\epsilon_i - |\epsilon| x_i)$$

Eigenfunction equations

$$\begin{aligned} \mathcal{L} Q_n^\circ(x) &= -\frac{1}{2} |n| (|n| + |\epsilon| - 1) Q_n^\circ(x) \\ \tilde{\mathcal{L}} \mathcal{D}(x, \epsilon) Q_n^\circ(x) &= -\frac{1}{2} |n| (|n| + |\epsilon| - 1) \mathcal{D}(x, \epsilon) Q_n^\circ(x) \end{aligned}$$

Mutant families in an infinite-leaf coalescent tree

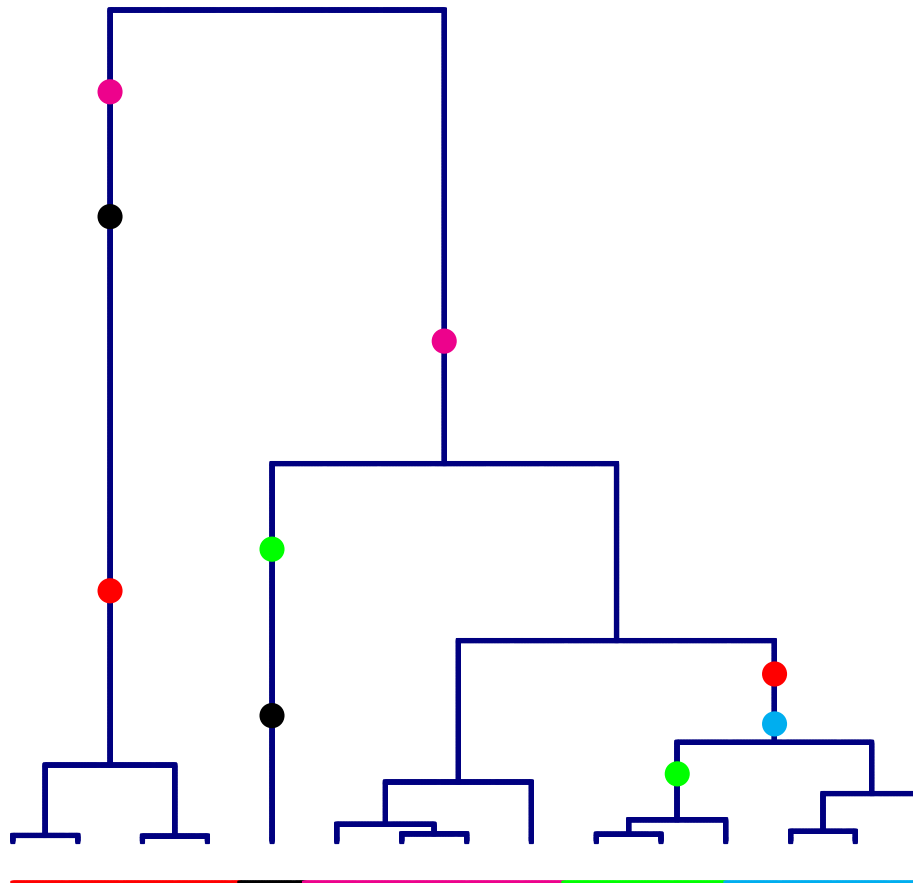
Mixture distribution arising from the coalescent

$$f(x, y; t) = \sum_{|l|=0}^{\infty} q_{|l|}^{|\epsilon|}(t) \sum_{\{l: |l| \text{ fixed}\}} \binom{|l|}{l} \prod_1^d x_i^{l_i} \mathcal{D}(y, \epsilon + l)$$

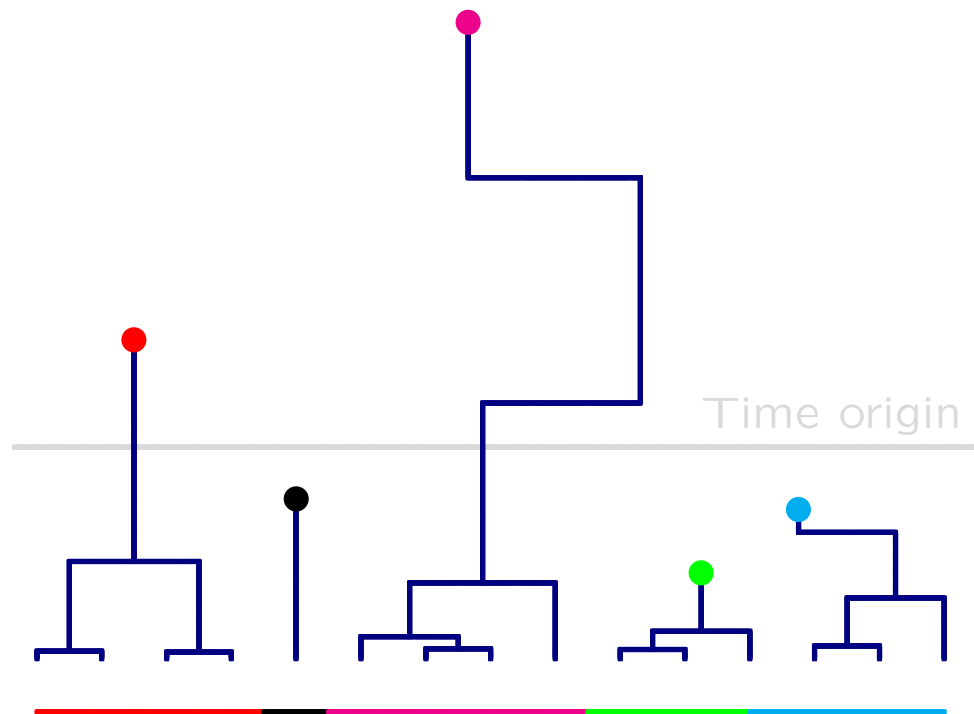
$q_{|l|}^{|\epsilon|}(t)$ is the distribution of $L^{|\epsilon|}(t)$, the number of non-mutant founder lineages at time t back. $L^{|\epsilon|}(t)$ is a death process back in time, starting from infinity, where lineages are reduced by coalescence or mutation from $k \rightarrow k-1$ at rate $k(k-1)/2 + k|\epsilon|/2$. Families are either from founder lineages or new mutations, giving the Dirichlet mixture.

$f(x, y; t)$ and the eigenfunction form are identical.

Infinite-leaf coalescent tree with mutations, type j with probability $\epsilon_j/|\epsilon|$



Forest of non-mutant ancestral lineages, to defining mutations



Ancestral lineages are **lost** back in time by coalescence or mutation at rates $\binom{i}{2}$ and $\frac{i\theta}{2}$ while i non-mutant lineages.

Dirichlet family sizes

$$\sum_{\{l: |l| \text{ fixed}\}} \binom{|l|}{l} \prod_1^d x_i^{l_i} \mathcal{D}(y, \epsilon + l)$$

$|l|$ non-mutant founder lineages are divided into $l = (l_1, \dots, l_d)$ numbers of types $1, \dots, d$ with probability $\binom{|l|}{l} \prod_1^d x_i^{l_i}$.

Let $U = (U_1, \dots, U_l)$ be their relative family sizes in the leaves of the tree, and $V = (V_1, \dots, V_d)$ be the frequencies of families derived from new mutations on the tree edges in $(0, t)$.

$U \oplus V = (U_1, \dots, U_l, V_1, \dots, V_d)$ is $\mathcal{D}(u \oplus v, (1, \dots, 1) \oplus \epsilon)$.

$\mathcal{D}(y, \epsilon + l)$ is obtained by adding Dirichlet parameters corresponding to types $1, \dots, d$.

Eigenfunction expansion

$$f(x, y; t) = \mathcal{D}(y, \epsilon) \left\{ 1 + \sum_1^{\infty} \rho_{|n|}(t) Q_{|n|}(x, y) \right\}$$

n -Kernel polynomials

$$Q_{|n|}(x, y) = \sum_{\{n; |n| \text{ fixed} \}} Q_n^{\circ}(x) Q_n^{\circ}(y)$$

invariant under which orthogonal polynomial set is used.

$$Q_{|n|}(x, y) = (|\epsilon| + 2|n| - 1) \sum_{m=0}^n (-1)^{|n|-m} \frac{(|\epsilon| + m)(|n|-1)}{m!(|n| - m)!} \xi_m$$

$$\text{where } \xi_m = \sum_{|l|=m} \binom{m}{l} \frac{|\epsilon|_{(m)}}{\prod_1^d \epsilon_i(l_i)} \prod_1^d (x_i y_i)^{l_i}$$

Two forms of the density

$$f(x, y; t) = \mathcal{D}(y, \epsilon) \left\{ 1 + \sum_1^{\infty} \rho_{|n|}(t) Q_{|n|}(x, y) \right\}$$

$\rho_{|n|}(t)$ are eigenvalues.

$$f(x, y; t) = \sum_{|l|=0}^{\infty} q_{|l|}^{|\epsilon|}(t) \sum_{\{l: |l| \text{ fixed} \}} \binom{|l|}{l} \prod_1^d x_i^{l_i} \mathcal{D}(y, \epsilon + l)$$

$q_{|l|}^{|\epsilon|}(t)$ are death process transition probabilities.

Fleming-Viot Dirichlet process version in Ethier & G (1993).

Poisson Dirichlet random measure

$$\mu = \sum_{i=1}^{\infty} x_i \delta_{\xi_i}$$

where $\{x_i\}$ is $\text{PD}(\theta)$ and independent of $\{\xi_j\}$ which are *i.i.d.* $\nu_0 \in \mathcal{P}(S)$, with S a compact metric space.

Stationary distribution of the random measure

$$\Pi_{\theta, \nu_0}(\cdot) = P(\mu \in \cdot)$$

Fleming-Viot process with type space S , and mutation operator

$$(Af)(x) = \frac{\theta}{2} \int_S (f(\xi) - f(x)) \nu_0(d\xi)$$

The **Fleming-Viot process** with type space S and mutation operator A has transition function $P(t, \mu, d\nu)$ for given $\mu \in \mathcal{P}(S)$

$$\begin{aligned}
 P(t, \mu, \cdot) = & q_0^\theta(t) \Pi_{\theta, \nu_0}(\cdot) \\
 & + \sum_{n=1}^{\infty} q_n^\theta(t) \int_{S^n} \mu^n(dy_1 \times \cdots \times dy_n) \\
 & \Pi_{n+\theta, (n+\theta)^{-1}\{n\eta_n(y_1, \dots, y_n) + \theta\nu_0\}}(\cdot)
 \end{aligned}$$

where $\eta_n(y_1, \dots, y_n)$ as the empirical measure of points $y_1, \dots, y_n \in S$,

$$\eta_n(y_1, \dots, y_n) = n^{-1}(\delta_{y_1} + \cdots + \delta_{y_n})$$

Characterization of Markov processes with
Beta stationary distributions and polynomial
eigenfunctions

Characterization of distributions with polynomial eigenfunctions

$$B(\alpha, \beta)^{-1} y^{\alpha-1} (1-y)^{\beta-1} \times \left\{ 1 + \sum_{n=1}^{\infty} \omega_n h_n R_n^{(\alpha, \beta)}(x) R_n^{(\alpha, \beta)}(y) \right\}, \quad 0 < x, y < 1$$

is a probability distribution for a sequence of constants $\{\omega_n\}$ if and only if $\{\omega_n\}$ is a positive definite sequence.

$\{R_n^{(\alpha, \beta)}(x)\}$ are Jacobi polynomials on the Beta distribution with $R_n^{(\alpha, \beta)}(1) = 1$. Scaling by $\sqrt{h_n}$ makes them orthonormal.

Bochner (1954) A bounded sequence $\{c_n\}$ is positive definite if

$$\sum a_n h_n R_n^{(\alpha, \beta)}(x) \geq 0, \quad \sum |a_n| h_n < \infty$$

implies that

$$\sum a_n c_n h_n R_n^{(\alpha, \beta)}(x) \geq 0$$

Gasper (1972) Let $\alpha < \beta$. If either $1/2 \leq \alpha$ or $\alpha + \beta \geq 2$, then a sequence ρ_n is positive definite if and only if

$$\rho_n = \mathbb{E}[R_n^{(\alpha, \beta)}(Z)]$$

for some random variable Z in $[0, 1]$.

Bochner (1954): A family of bounded sequences $\{c_n(t)\}$, $0 \leq t < \infty$ is called a homogeneous stochastic process if

- (i) $\{c_n(t)\}$ is positive definite for each t
- (ii) $c_n(t)$ is continuous in t
- (iii) $c_n(0) = c_0(t) = 1$
- (iv) $c_n(t + s) = c_n(t)c_n(s)$

Gaspar (1972): Under the conditions on α and β , it is N & S that $c_n(t) = e^{-d_n t}$ with

$$d_n = \sigma n(n + \alpha + \beta - 1) + \int_0^1 \frac{1 - R_n^{(\alpha, \beta)}(z)}{1 - z} d\nu(z)$$

where $\sigma \geq 0$ and ν is a finite measure on $[0, 1)$.

Sequences $\{c_n(t)\}$ characterize eigenvalues of transition probability functions with polynomials as eigenfunctions.

Subordinated Diffusion Process Let $\{X(t)\}$ be the two allele Wright-Fisher diffusion process with eigenvalues $e^{-\frac{1}{2}n(n+\theta-1)t}$, where $\theta = \alpha + \beta$

Let $\{Z(t)\}$ be a Lévy process with Laplace transform

$$\mathbb{E}[\exp(-\phi Z(t))] = \exp \left\{ -t \int_0^\infty \frac{1 - e^{-y\phi}}{y} G(dy) \right\}$$

where $\int_1^\infty \frac{1}{y} G(dy) < \infty$.

Then $\{X(Z(t))\}$ has polynomial eigenvectors and eigenvalues

$$c_n(t) = \exp \left\{ -t \int_0^\infty \frac{1 - e^{-yn(n+\theta-1)/2}}{y} G(dy) \right\}$$

Markov Process $\{\widetilde{X}(t)\}$ with transition density

$$B(\alpha, \beta)^{-1} y^{\alpha-1} (1-y)^{\beta-1} \\ \times \left\{ 1 + \sum_{n=1}^{\infty} e^{-nt} h_n R_n^{(\alpha, \beta)}(x) R_n^{(\alpha, \beta)}(y) \right\}, \quad 0 < x, y < 1$$

Poisson kernel in orthogonal polynomial theory

$$1 + \sum_{n=1}^{\infty} r^n h_n R_n^{(\alpha, \beta)}(x) R_n^{(\alpha, \beta)}(y), \quad 0 < x, y < 1, |r| < 1$$

$\widetilde{X}(t) = X(Z(t))$, where $\{Z(t)\}$ is a Lévy process with Laplace transform

$$\exp \left\{ -t \left[\sqrt{2\lambda + (\theta - 1)^2/4} - \sqrt{(\theta - 1)^2/4} \right] \right\}, \quad \theta \geq 1$$

$Z(t)$ is a tilted positive stable process with index $\frac{1}{2}$ and density

$$\frac{t}{2\pi z^3} e^{-\frac{1}{2}\frac{t}{\sqrt{z}}} \cdot e^{-\frac{1}{8}(\theta-1)^2 z + \frac{1}{2}|(\theta-1)|t}, \quad z > 0$$

$Z(t)$ is a Lévy process with Laplace transform

$$\exp \left\{ -\frac{t}{\sqrt{2\pi}} \int_0^\infty \frac{e^{-x(\theta-1)^2/8}}{x^{3/2}} (1 - e^{-x\lambda}) dx \right\}$$

The eigenvalues of $\widetilde{X}(t)$ are, for $\theta \geq 1$,

$$\mathbb{E} \left[\exp \left\{ -\frac{1}{2}n(n + \theta - 1)Z(t) \right\} \right] = \exp\{-nt\}$$

$\widetilde{X}(t)$ has a generator \widetilde{L} satisfying

$$-\widetilde{L} = \sqrt{2(-L) + (\theta - 1)^2/4} - \sqrt{(\theta - 1)^2/4}$$

A series expansion with positive coefficients of L is

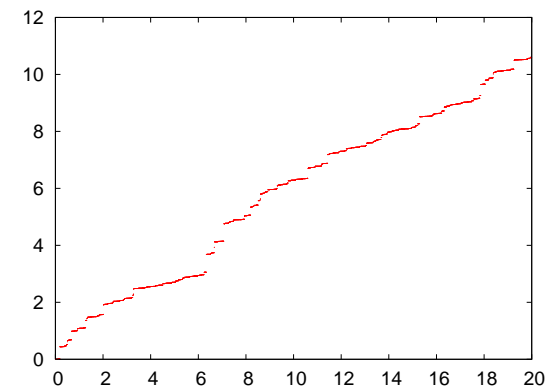
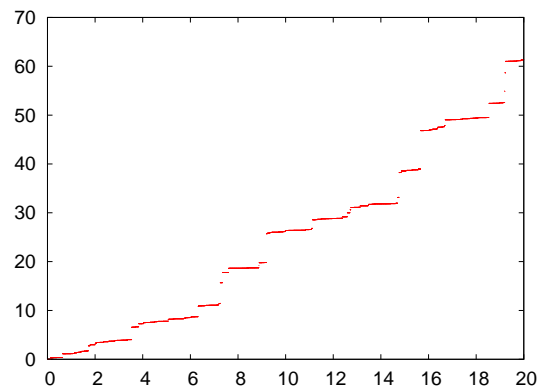
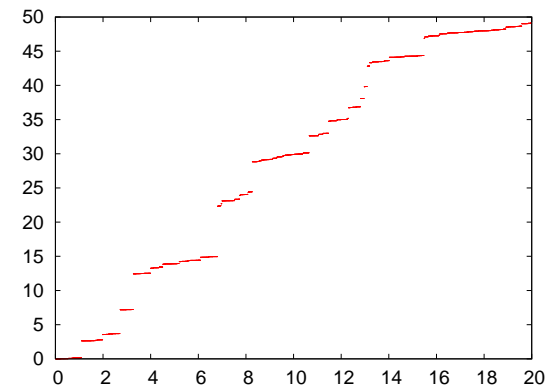
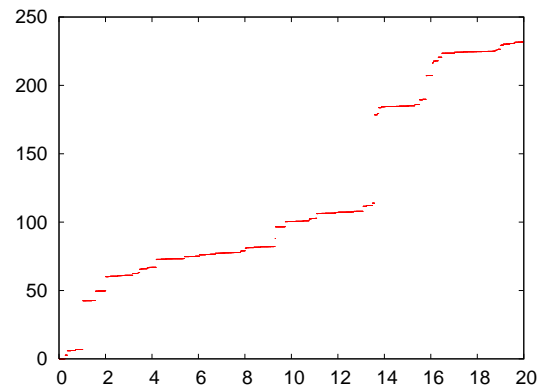
$$\widetilde{L} = \sqrt{c} \sum_{k=1}^{\infty} \frac{2^k}{k!} \left(\frac{1}{2}\right)_{(k-1)} \left(\frac{L}{c}\right)^k$$

where $c = (\theta - 1)^2/4$.

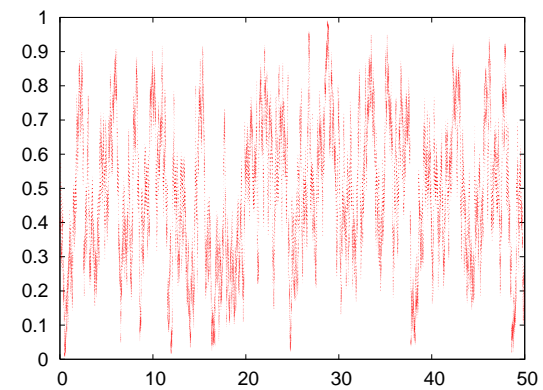
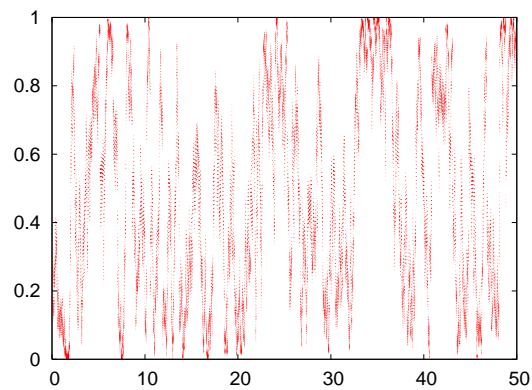
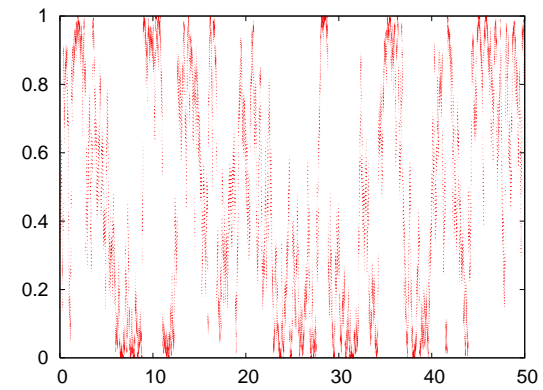
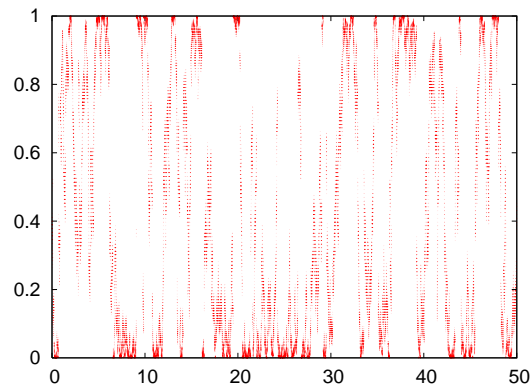
$$L = \frac{1}{2} \left(-\widetilde{L}^2 + (\theta - 1)\widetilde{L} \right)$$

Eigenfunctions

$$\widetilde{L} R_n^{(\alpha, \beta)}(x) = -n R_n^{(\alpha, \beta)}(x)$$



Tilted Stable subordinator, index half ($\theta = 1.0, 1.5, 2.0, 5.0$)



Subordinated Wright-Fisher diffusion ($\theta = 1.0, 1.5, 2.0, 5.0$)

If $\theta < 1$ then for $n \geq 1$,

$$\mathbb{E}\left[\exp\left\{-\frac{1}{2}n(n+\theta-1)Z(t)\right\}\right] = \exp\{-nt\} \times \exp\{t(1-\theta)\}$$

Let $\tilde{f}(x, y; t)$ be the transition density of $\tilde{X}(t)$, then the transition density with eigenvalues $\exp\{-nt\}$, $n \geq 0$ is

$$e^{-t(1-\theta)}\tilde{f}(x, y; t) + (1 - e^{-t(1-\theta)})B(\alpha, \beta)^{-1}y^{\alpha-1}(1-y)^{\beta-1}$$

Subordinated process $X(\hat{Z}(t))$ where $\hat{Z}(t)$ is a similar process to $Z(t)$ but has an extra state ∞ . $Z(t)$ is killed by a jump to ∞ at a rate $(1-\theta)$.

Subordinated genealogical form of the transition density

$$\sum_{\ell=1}^{\infty} q_{\ell}^{\theta}(t) \sum_{k=1}^{\ell} \binom{\ell}{k} p^k (1-p)^{\ell-k} B(\alpha+k, \beta+\ell-k)^{-1} x^{\alpha+k-1} (1-x)^{\beta+\ell-k-1}$$

where $\{q_{\ell}^{\theta}(t)\}$ are transition probabilities of a death process $\{A^{\theta}(t), t \geq 0\}$ with death rates $\mu_k = \binom{k}{2} + \frac{k\theta}{2}$ and $A^{\theta}(0) = \infty$.

Subordinated process $\widetilde{X}(t) = X(Z(t))$ has a similar form for the transition density, with $q_{\ell}^{\theta}(t)$ replaced by $\mathbb{E}(q_l^{\theta}(Z(t)))$, transition functions of the subordinated death process $A^{\theta}(Z(t))$.

Subordinated death process $A^\theta(Z(t))$ example

$Z(t)$ is a tilted positive stable process with index $\frac{1}{2}$.

$\tilde{A}^\theta(t) = A^\theta(Z(t))$ is a Markov process with $\tilde{A}^\theta(0) = \infty$ and a pgf of

$$G_{\tilde{A}^\theta(t)}(s) = \left(\frac{1 - 4pq s}{1 - 4pq} \right)^{-\frac{1}{2}} \times \left(\frac{1 - \sqrt{1 - 4pq s}}{2ps} \right)^{\theta-1}, \quad \theta \geq 0,$$

where $p = e^{-t}/(1 + e^{-t})$ and $q = 1/(1 + e^{-t})$.

Random walk on \mathbb{Z} with transitions $j \rightarrow j + 1$ with probability p and $j \rightarrow j - 1$ with probability $q = 1 - p$ and $q \geq p$.

Let the number of steps to hit $-\theta$, starting from 0 be $2\xi + \theta$.

Then ξ has a pgf of

$$H(s) = \left(\frac{1 - \sqrt{1 - 4pqs}}{2ps} \right)^\theta$$

$\tilde{A}(t) + 1$ has the same distribution as the size-biased distribution of ξ , with pgf

$$G_{\tilde{A}(t)} = \frac{sH'(s)}{H'(1)}$$

The probability distribution of $\tilde{A}^\theta(t)$ is

$$\binom{2j + \theta - 1}{j} \left(\frac{z}{1+z}\right)^j \left(\frac{1}{1+z}\right)^{j+\theta} (1-z)$$

for $j = 0, 1, \dots$, where $z = e^{-t}$.

The probability distribution of $A^\theta(t)$ is

$$\sum_{k=j}^{\infty} \rho_k^\theta(t) (-1)^{k-j} \frac{(2k + \theta - 1)(j + \theta)_{(k-1)}}{j!(k-j)!}$$

for $j = 0, 1, \dots$, where $\rho_k^\theta(t) = e^{-k(k+\theta-1)t/2}$.

General characterization of eigenvalues

$$R_n^{(\alpha, \beta)}(x) = {}_2F_1(-n, n + \theta - 1; \alpha; 1 - y)$$

Eigenvalues

$$\begin{aligned} & \int_0^{1-} \frac{1 - R_n^{(\alpha, \beta)}(y)}{1 - y} dG(y) \\ &= -c \sum_{k=1}^n \frac{(-n)_{(k)}(n + \theta - 1)_{(k)} \mu_{k-1}}{\alpha_{(k)} k!} \\ &= -c \sum_{k=1}^n \frac{\prod_{j=0}^{k-1} (-n(n + \theta - 1) + j(j + \theta - 1)) \mu_{k-1}}{\alpha_k k!} \end{aligned}$$

where $\int_0^{1-} (1 - y)^k dG(y) = c\mu_k$.

The generator corresponding to a process with these eigenvalues is

$$\hat{L} = c \sum_{k=1}^{\infty} \frac{\prod_{j=0}^{k-1} (2L + j(j + \theta - 1))}{\alpha_{(k)}} \frac{\mu_{k-1}}{k!}$$

This is a positive series in L .

Laguerre diffusion $\{Y(t), t \geq 0\}$

Branching process with immigration

Generator

$$L = x \frac{\partial^2}{\partial x^2} + (-bx + c) \frac{\partial}{\partial x}, \quad b = 1, c > 0$$

Transition functions, for $x, y > 0$

$$\begin{aligned} f(x, y; t) &= \frac{y^{c-1}}{\Gamma(c)} e^{-y} \sum_{n=0}^{\infty} e^{-nt} L_n^{(c-1)}(x) L_n^{(c-1)}(y) \frac{n!}{c(n)} \\ &= \sum_{k=0}^{\infty} e^{-\mu x} \frac{(\mu x)^k}{k!} \frac{(1-r)^{-(c+k)}}{\Gamma(c+k)} y^{c+k-1} e^{-y/(1-r)} \end{aligned}$$

where $\{L_n^{(c-1)}(x)\}$ are Laguerre polynomials, $r = e^{-t}$ and $\mu = r/(1-r)$.

Transition functions of a Markov process

$$\tilde{f}(x, y; t) = \frac{y^{c-1}}{\Gamma(c)} e^{-y} \sum_{n=0}^{\infty} e^{-td_n} L_n^{(c-1)}(x) L_n^{(c-1)}(y) \frac{n!}{c_{(n)}}$$

are characterized by their eigenvalues having the representation

$$d_n = \int_0^1 \frac{1 - y^n}{1 - y} dG(y) = an + \int_0^{1-} \frac{1 - y^n}{1 - y} dG(y), \quad a \geq 0$$

Sarmanov (1968), G (1969).

A Markov process with transition functions $\tilde{f}(x, y; t)$ has a characterization as a subordinated process

$$\tilde{Y}(t) = Y(at + Z(t))$$

where $\{Z(t), t \geq 0\}$ is a Lévy process with Laplace transform

$$\exp \left\{ - \int_0^{1-} \frac{1 - e^{-\lambda(-\log y)}}{1 - y} dG(y) \right\}$$

$$\int_0^{1-} 1 \wedge (-\log y) \frac{1}{1 - y} dG(y) < \infty \equiv \int_0^{1-} dG(y) < \infty$$

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